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Spontaneous Subarachnoid Hemorrhage: A Systematic Review and Meta-Analysis Describing the Diagnostic Accuracy of History, Physical Exam, Imaging, and Lumbar Puncture with an Exploration of Test Thresholds

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#### **Abstract**

**Background:** Spontaneous subarachnoid hemorrhage (SAH) is a rare, but serious etiology of headache. The diagnosis of SAH is especially challenging in alert, neurologically intact patients, as missed or delayed diagnosis can be catastrophic.

Objectives: The objective was to perform a diagnostic accuracy systematic review and metaanalysis of history, physical examination, cerebrospinal fluid (CSF) tests, computed
tomography (CT), and clinical decision rules for spontaneous SAH. A secondary objective
was to delineate probability of disease thresholds for imaging and lumbar puncture (LP).

Methods: PubMed, Embase, Scopus, and research meeting abstracts were searched up to
June 2015 for studies of emergency department patients with acute headache clinically
concerning for spontaneous SAH. QUADAS-2 was used to assess study quality and, when
appropriate, meta-analysis was conducted using random effects models. Outcomes were
sensitivity, specificity, and positive (LR+) and negative (LR-) likelihood ratios. To identify
test and treatment thresholds, we employed the Pauker-Kassirer method with Bernstein test
indication curves using the summary estimates of diagnostic accuracy.

*Results:* A total of 5,022 publications were identified, of which 122 underwent full-text review; 22 studies were included (average SAH prevalence = 7.5%). Diagnostic studies differed in assessment of history and physical examination findings, CT technology, analytical techniques used to identify xanthochromia, and criterion standards for SAH. Study quality by QUADAS-2 was variable; however, most had a relatively low risk of biases. A history of neck pain (LR+ = 4.1; 95% confidence interval [CI] = 2.2 to 7.6) and neck stiffness on physical examination (LR+ = 6.6; 95% CI = 4.0 to 11.0) were the individual findings most strongly associated with SAH. Combinations of findings may rule out SAH, yet promising clinical decision rules await external validation. Noncontrast cranial CT within 6 hours of headache onset accurately ruled in (LR+ = 230; 95% CI = 6 to 8,700) and ruled out SAH

(LR-= 0.01; 95% CI = 0 to 0.04); CT beyond 6 hours had a LR- of 0.07 (95% CI = 0.01 to 0.61). CSF analyses had lower diagnostic accuracy, whether using red blood cell (RBC) count or xanthochromia. At a threshold RBC count of  $1,000 \times 10^6$ /L, the LR+ was 5.7 (95% CI = 1.4 to 23) and LR- was 0.21 (95% CI = 0.03 to 1.7). Using the pooled estimates of diagnostic accuracy and testing risks and benefits, we estimate that LP only benefits CT-negative patients when the pre-LP probability of SAH is on the order of 5%, which corresponds to a pre-CT probability greater than 20%.

Conclusions: Less than one in 10 headache patients concerning for SAH are ultimately diagnosed with SAH in recent studies. While certain symptoms and signs increase or decrease the likelihood of SAH, no single characteristic is sufficient to rule in or rule out SAH. Within 6 hours of symptom onset, noncontrast cranial CT is highly accurate, while a negative CT beyond 6 hours substantially reduces the likelihood of SAH. LP appears to benefit relatively few patients within a narrow pretest probability range. With improvements in CT technology and an expanding body of evidence, test thresholds for LP may become more precise, obviating the need for a post-CT LP in more acute headache patients. Existing SAH clinical decision rules await external validation, but offer the potential to identify subsets most likely to benefit from post-CT LP, angiography, or no further testing.

# Introduction

Headache is a common presenting complaint to the Emergency Department (ED), and represents approximately 2% of ED visits annually in the United States. While most headaches are self-limiting, the possibility of spontaneous subarachnoid hemorrhage (SAH) is an important diagnostic consideration in patients presenting with a sudden, severe headache. These sudden onset headaches include the classic "thunderclap headache" that instantaneously peaks at headache onset, as well as headaches that reach maximal severity within seconds to one hour of onset. Sudden onset headaches include a wide spectrum of possible diagnoses, including SAH (approximately 1% of acute

headaches), benign post-coital headache, exertional headache, intracranial cysts or tumors, intracerebral hemorrhage, hypophyseal apoplexy, sphenoid sinusitis, sinus thrombosis, cough-related, vascular dissection, cerebral vasospasm, and migraine headaches. High more benign headaches mimic SAH and are estimated to be at least 50-times more common than SAH. Moreover, while SAH is often caused by a ruptured cerebral aneurysm and represents a neurosurgical emergency, a variety of other SAH etiologies range from the benign, low-pressure perimesencephalic hemorrhage which can be treated conservatively 12,13 to vascular malformations, arterial dissection, and vasculitis which require time-sensitive interventions (Table 1). Thus, the diagnostic approach to acute headache epitomizes the practice of emergency medicine with a high stakes condition without a clear-cut presentation lurking within a high volume complaint, and ultimately most patients do not have a serious diagnosis.

Between 1960 and 1995, the six-month mortality of aneurysmal SAH decreased by 15% <sup>15</sup>, but mortality remains 27%-44% with substantial regional variability. <sup>16</sup> Up to 25% of patients die within 24 hours, and the three-month mortality rate can be as high as 50% without early definitive treatment. <sup>17</sup> In addition, one-third of SAH survivors suffer neurological deficits affecting their daily lives <sup>15</sup> and up to 50% never return to work. <sup>18</sup> Early detection and treatment can significantly reduce morbidity and mortality. <sup>11,19</sup> However, SAH can be a challenging diagnosis to make, particularly in alert, neurologically intact adults. <sup>20</sup> Missed diagnosis is the primary reason for delayed treatment and case series suggest that 12%-53% of SAH cases are misdiagnosed on their initial presentation in a variety of settings, including the ED. <sup>21</sup> Patients with misdiagnosed or undiagnosed warning bleeds that subsequently re-bleed have a 70% mortality. <sup>22-25</sup> ED providers miss the diagnosis of SAH in about 5% of cases, primarily those with lower acuity at presentation and misdiagnosis is estimated to be more likely in non-teaching hospitals. <sup>26,27</sup> Most emergency medicine and neurosurgery textbooks provide limited guidance on the diagnostic accuracy and utility of history, physical exam, cerebrospinal fluid (CSF) analysis, or CT for the diagnosis of SAH. <sup>28,29</sup> Indeed, given the serious nature of the condition, few have proposed a sensible threshold below which testing is not necessary,

which contributes to over-testing and the unintended consequences that include incidental findings with prolonged patient uncertainty.<sup>30</sup>

Unenhanced cranial CT is usually the first-line diagnostic test to evaluate for suspected SAH. Because early generation CTs did not detect up to 5% of SAHs, traditional teaching has emphasized the need to perform lumbar puncture (LP) to exclude SAH to reduce the probability of error. Yet many patients fear having an LP because an LP is painful, invasive, and can lead to a post-LP headache in 6%-30% of patients depending on the gauge of the LP needle.<sup>31</sup> Additional clinician-level barriers to LP testing include the time needed to perform<sup>32</sup> and low diagnostic yield.<sup>20,33-36</sup> In addition, LP may be contraindicated if a patient has a bleeding disorder, increased intracranial pressure, or simply does not consent to the procedure.<sup>37</sup> Despite dogma, LP is not always performed after CT and multicenter observational studies report that fewer than half of acute headache patients undergo LP following a negative CT to "rule-out" SAH. Even in these studies less than one in 100 LPs are ultimately deemed "true positives" after negative CT using a third generation scanner. 38,39 Additionally, the specificity of LP is reduced by "traumatic taps" in which bleeding from the procedure contaminates the CSF being collected, estimated to occur in about one in six LPs. 40 Identifying xanthochromia due to the breakdown of hemoglobin in CSF is considered to be pathognomonic for SAH, yet xanthochromia requires up to 12 hours to develop, and is prone to false positives due to ex vivo hemolysis or hyperbilirubinemia. 21,41 National guidelines in the United Kingdom advocate using spectrophotometry to identify xanthochromia, yet nearly all North American hospitals rely on visual inspection alone. 42 A recent systematic review comparing visual inspection to spectrophotometry to detect CSF xanthochromia was unable to draw a conclusion due to significant heterogeneity in the underlying studies.43

To our knowledge, there are no prior systematic reviews assessing the diagnostic accuracy of history, physical exam, and imaging studies for SAH in the setting of acute onset headache patients in the ED.

Therefore, we set out to collect and summarize the available published literature regarding these characteristics as well as LP findings and clinical decision rules for spontaneous (non-traumatic) SAH. Since recent systematic reviews have evaluated CT angiography (CTA)<sup>44</sup> and magnetic resonance angiography (MRA),<sup>45</sup> we did not repeat these meta-analyses. However, additional studies exploring visible xanthochromia diagnostic accuracy have been published since the 2014 systematic review by Chu et al<sup>43</sup> so we report an updated meta-analysis of visible xanthochromia. We also used these diagnostic accuracy estimates to delineate an optimal diagnostic strategy using the test-treatment threshold approach of Pauker and Kassirer. In particular, we used the graphical Bayesian approach of test-indication curves first proposed by Bernstein<sup>48</sup> to identify the *a priori* likelihood of disease in patients for whom LP would be rational despite a negative cranial CT.

#### Methods

Study Design

We conducted a systematic review and meta-analysis of original research studies that reported data on the diagnosis of SAH in ED patients with acute headache. The design and manuscript structure conform to the recommendations from the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA)<sup>49</sup> and the Meta-analysis of Observational Studies (MOOSE) criteria.<sup>50</sup> Studies were included if they described adult patients (14 years or older) seen in an ED with acute headache or other symptoms or signs (such as syncope, acute mental status change, or otherwise unexplained nausea) and in whom SAH was a diagnostic consideration. In addition, studies had to report sufficient detail on the diagnostic test and criterion standard to be able to construct two-by-two tables. We elected to define "disease positive" as being an acute SAH including non-aneurysmal SAH (e.g. perimesencepablic hemorrhage) as most original studies did not distinguish between SAH subtypes. We recognized that CT findings of subarachnoid blood were therefore incorporated into this definition, thereby inflating CT sensitivity, specificity, LR+ and LR- (incorporation bias), but report these estimates for completeness.<sup>51</sup> Recognizing the absence of an accepted definition for a

positive LP, we analyzed separately an erythrocyte count above  $1,000 \times 10^6 / L$  in the final tube as well as either visible or spectrophotometric xanthochromia. Patients with negative imaging in whom LP was not performed were deemed as "disease negative" only when the authors reported effort at follow-up after hospital discharge. We excluded studies that solely reported data on patients with SAH (i.e. where only sensitivity could be calculated), precluding an estimate of specificity or likelihood ratios.

Search Strategy

We searched the published literature using strategies created for the concepts of Emergency

Department, Headache, Subarachnoid Hemorrhage, Physical Exam, and Diagnostic Accuracy. Search

strategies were established using a combination of standardized terms and key words, and were

implemented in PubMed 1946-, Embase 1947-, and Scopus 1823-. All searches were completed in

June 2015 and limited to English using database supplied limits. Full search strategies for PubMed

and Embase are provided in the Online Appendix 1. One author (AMH) independently reviewed the

titles and abstracts to identify potentially relevant articles. Where there were questions about

inclusion, this was discussed with a second author (JMP) and a consensus decision was made about

whether the article should be included for full-text review. Studies deemed not to contain clear

inclusion or exclusion criteria were evaluated by a second author (JMP) before final inclusion or

exclusion.

Data from each study were independently abstracted by two authors (AMH, CRC). Information abstracted included the individual study setting, inclusion criteria, exclusion criteria, specific CT (machine, slice number, and type of CT image), definition of disease, criteria for a positive LP, follow-up description, and whether or not there was neurosurgical intervention. The abstracted data were constructed into tables, which were confirmed by a second author (CRC); any uncertainty was discussed between the two authors and a consensus decision was made. In addition, the references for each full-text article were reviewed to assess whether additional studies could be included in the review, which were then assessed by full-text review. The authors conducted bibliographic searches

of research abstract presented at scientific meetings published in *Academic Emergency Medicine*, Annals of Emergency Medicine, Stroke, Neurology, and Neurosurgery, from 2002 through October 2015.

Individual Evidence Quality Appraisal

Two authors (AMH, MJW) independently used the Quality Assessment Tool for Diagnostic Accuracy Studies (QUADAS-2)<sup>52</sup> for systematic reviews to evaluate the quality of evidence for the studies identified. These two authors used several *a priori* conditions to evaluate an individual study's risk of bias and degree of applicability.

- In the assessment of "inappropriate exclusions" reviewers assessed whether a study's inclusion and exclusion criteria led investigators to evaluate suspected SAH patients who were more or less acutely ill than those typically evaluated in ED settings, which could introduce spectrum bias<sup>51</sup> or spectrum effect<sup>53</sup> and skew observed estimates of sensitivity (with sicker populations) or specificity (with less sick populations) upwards.
- If the study setting was anything other than an ED (for example, a neurosurgery clinic) then the results were assessed as "low applicability".
- In the assessment of continuous data like CSF RBCs, if the authors pre-defined a threshold for abnormal (example, CSF RBC >  $5 \times 10^6$  cells/L in the last tube is "abnormal") then the study results were assessed as "low risk" of bias.
- If the conduct of the index test differed from routine care (example, xanthochromia assessed using only spectrophotometry), then the results were assessed as "low applicability".
- No widely accepted criterion standard for SAH yet exists, so we accepted the gold standard of subarachnoid blood on unenhanced CT of the head, or CSF xanthochromia, or CSF RBCs > 5 x 10<sup>6</sup>/L in the final sample of CSF when any of these findings included an aneurysm or arteriovenous malformation evident on cerebral angiography, as previously established by a consensus panel of emergency physicians and a neurosurgeon.<sup>39,54</sup> If less than 100% of a

study's subjects underwent cerebral angiography, a surrogate outcome of medical record review and structured telephone follow-up of at least 6-months was necessary.<sup>20</sup> Any other criterion standard was assessed as "high risk" for bias.

- Assessing whether the target condition is identified is challenging in SAH diagnostic research. The ideal patient-centric diagnostic test would identify SAH patients with a lesion amenable to surgical or endovascular intervention rendering measurable benefit in mortality and morbidity. Since perimesencephalic bleeds cause about 10% of SAH cases yet require no intervention other than symptom control, diagnostic studies that report aneurysmal SAH separately from perimesencephalic SAH would provide a higher level of evidence by which to associate testing with patient-benefit. Unfortunately, few SAH studies report aneurysmal versus non-aneurysmal SAH separately, so the decision was made that any SAH identified using an acceptable criterion standard as defined above would be assessed as "high applicability".
- In assessing the uniformity of the criterion standard reported, if all patients with non-diagnostic CT and LP had either subsequent advanced imaging (angiography) or post-ED follow-up, then the risk of differential verification bias was assessed as being "low".<sup>51</sup>
- If a substantial number of patients were reported lost to follow-up and no sensitivity analysis assuming worst-case scenario (lost to follow-up = false negative SAH index test), then the risk of bias was assessed to be "high".

Inter-rater QUADAS-2 reliability was assessed using prevalence adjusted, bias adjusted kappa values to adjust for unique biases between reviewers and the distribution of responses across categories.<sup>56</sup> Discrepancies were then resolved by consensus.

#### Data Collection

We abstracted the study setting, study inclusion criteria, the criterion standard employed to diagnose SAH, CT technology, CSF analysis, follow-up procedures in patients who did not undergo both CT and LP, disease prevalence, and every variable on history, physical, or testing for which a 2x2 diagnostic accuracy table could be constructed. When reconstructed 2x2 contingency tables differed from the original investigators' reports of sensitivity and specificity, we contacted the authors for clarity.

Data Analysis

Meta-analysis estimates were computed by one author (CRC) when ≥1 studies assessed the same finding on history or physical examination, or index testing. We chose to combine unenhanced cranial CT studies without distinguishing specific generations of technology since the detector number or geometry and the thickness of slices were frequently not specified. We generated combined estimates for diagnostic accuracy using a random-effects model (Meta-DiSc Version 1.4, Hospital Universitario Ramón y Cajal, Madrid, Spain). The Interstudy heterogeneity was assessed using the Der-Simonian-Laird random effects model and the Index of Inconsistency (I²). Pooled estimates of dichotomous positive (LR+) and negative (LR-) likelihood ratios are also reported from the random effects model. Publication bias was not assessed because this is not an accepted approach in diagnostic meta-analyses due to the small number of studies generally identified.

The added value of likelihood ratios over positive and negative predictive values is that likelihood ratios are independent of disease prevalence. <sup>62</sup> In addition, likelihood ratios can be used on an individual patient to estimate post-test probabilities, unlike sensitivity and specificity. <sup>63</sup> Larger LR+ increase the post-test probability of a disease more than smaller values of LR+; similarly, smaller LR-decrease the post-test probability of a disease more than larger values of LR-. Note that a likelihood

ratio of 1 does not change the post-test probability at all. One rule of thumb is that LR+>10 and LR-<0.1 provide considerable diagnostic value.<sup>64</sup>

# Test-Treatment Threshold

We used the Pauker-Kassirer method to estimate thresholds for further testing or treatment.<sup>47</sup> This technique is based on the diagnostic accuracy of a test, as well as the estimated risks of testing and of treatment, and the anticipated benefit of treatment. We performed a separate analysis for each of the two primary tests in question, namely cranial CT (stratified by time post headache onset) and LP (using either CSF RBC or visible xanthochromia to define "positive LP"). We used the pooled estimates from the meta-analysis for test sensitivity and specificity, and our collective clinical experience and selected references to estimate risks and benefits. We extended the test-indication curve method first proposed by Bernstein<sup>48</sup> to visually represent this analysis. Specifically, the curves display how the information conveyed by a positive or negative test result affects a given patient's post-test probability of disease relative to the treatment threshold. These test-indication curves were originally developed to illustrate the pre-test probability at which testing is no longer indicated based only on the net utilities of treating or not treating patients with and without disease. 46 We extended this graphical approach to incorporate the non-zero risks of testing, which necessarily narrow the range of probabilities for which testing remains rational. 47,65 Because the precise utilities of both testing and treatment are difficult to estimate on a common scale, we performed sensitivity analyses by doubling and halving the calculated treatment threshold (i.e. the ratio of risks of treatment to the sum of risks and benefits), 48 and also by reducing the risk of testing by half and recalculating the test thresholds. Recognizing that these estimates are nevertheless somewhat arbitrary, we also provide an interactive calculator to permit readers to re-compute thresholds using different estimates of test performance or anticipated risks and benefits that may be more applicable to the end users' patient populations and clinical environments.

In developing the test-treatment threshold, we were most interested in the common clinical dilemma of whether to perform LP after a negative cranial CT, or to forgo LP and proceed directly to angiography. Because CT-positive or LP-positive patients do not immediately undergo neurosurgical clipping or endovascular coiling, the next step being contemplated is usually angiography, typically by CT or MRA (and sometimes by digital subtraction fluoroscopy), followed by intervention when a culprit aneurysm is identified. Therefore, for Pauker-Kassirer modeling we designated the "test" of interest as being either the unenhanced CT or the LP, and the "treatment" as being angiography. Furthermore, because an LP is usually done after a negative CT, we considered how a negative unenhanced cranial CT would affect the post-CT (i.e. pre-LP) probability of still having a SAH across the range of pre-CT probabilities.

#### **Results**

A total of 5,022 citations were identified through a search of PUBMED, EMBASE, and SCOPUS. From this search, 399 duplicate citations were removed. 122 studies underwent full text-review; of these, 22 studies were included in the current analysis. (Figure 1 and Table 2). Although most studies enrolled only awake and alert patients with a GCS of 15, some also included those with decreased level of consciousness which could introduce spectrum effect into observed estimates of diagnostic accuracy. All studies were conducted in North American or European countries. Each study included the presence of subarachnoid blood on unenhanced cranial CT as representing "disease positive".

Descriptors of the specific technology utilized for CT scanners and spectrophotometric assessment for CSF xanthochromia were inconsistent. Definitions of xanthochromia ranged from direct visual inspection to spectrophotometric analysis. When using spectrophotometry, studies varied with regards to wavelengths sampled and specific algorithms used to adjust for interference by oxyhemoglobin, protein, and other corrections. Clinical follow-up proportion varied considerably, ranging from no follow-up to complete follow-up with cross-checking of coroner records to ensure no missed deaths. Assessment of studies by QUADAS-2 (Table 3) showed significant heterogeneity in

reference standards used, blinding of index test interpreters to the reference standard and vice versa. Agreement between reviewers was fair for all QUADAS-2 elements except for index and reference standard blinding, reference standard appropriateness, and acceptable application of index test in which cases the agreement was poor. However, most studies had a relatively low-risk of biases overall.

#### Prevalence

The overall prevalence of SAH in these studies ranged from 0.91% to 68%. In prospective studies within the last ten years, the prevalence ranged from 0.91% to 17% with a weighted average of 7.5%. Two of the studies with the lowest prevalence enrolled patients undergoing LP following a negative CT. <sup>67,68</sup>

# History and Physical Examination

Eight studies described the diagnostic accuracy for 22 components of history  $^{2,17,33,38,39,69-71}$  (Table 4 and Figure 2) and six studies described the diagnostic accuracy for four physical exam tests  $^{2,17,69,70}$  (Table 4 and Figure 3) for SAH. Within these studies there was significant heterogeneity in the manner in which history and physical characteristics were obtained and defined. For example, Perry et al. defined limited neck flexion on physical exam as "inability to touch chin to chest or raise the head 8 cm off the bed if supine",  $^{39}$  whereas Carstairs provided no descriptor of how "nuchal rigidity" was objectively assessed on physical exam.  $^{17}$  Perry et al. reported the inter-physician reliability of elements of the history and physical exam, with kappa ( $\kappa$ ) values ranging from 0.44 to 1.00. The least reliable findings were thunderclap headache ( $\kappa$ = 0.49) and limited neck flexion ( $\kappa$ = 0.44).  $^{39}$ 

No single element of history had a very high pooled LR+. The findings on history that increased the probability of SAH the most included subjective neck stiffness (LR+ 4.12, 95% CI 2.24-7.59), while the absence of "worst headache of life" (LR- 0.36, 95% CI 0.01-14.22) or onset of headache over more than one hour (LR- 0.06, 95% CI 0-0.95) each reduced the probability of SAH.

On physical exam, neck stiffness (LR+ 6.59, 95% CI 3.95-11.00) was strongly associated with SAH as reported in three studies, but no single physical exam finding had a LR- less than 0.74. Most elements of history and physical exam demonstrated significant statistical heterogeneity.

#### Clinical Decision Rules

Four related SAH clinical decision rules have been described, three of which were prospectively validated in a subsequent study by the same investigator group to generate the final refined Ottawa SAH Rule. Since more than one study is needed to perform a meta-analysis, we were unable to pool diagnostic accuracy results for SAH Clinical Decision Rules. The component elements of each decision rule, along with their combined accuracy and reliability are available in Online Appendix 2. Rule 1 appears sufficient to rule-out SAH (LR- 0.06, 95% CI 0.01-0.22), was uncomfortable to use for only 18% of surveyed emergency physicians, was misinterpreted in 4.7% of cases, and would theoretically decrease CT and/or LP testing rates from 84% to 74%. On the other hand, the Ottawa SAH Rule more accurately rules out SAH (LR- 0.02, 95% CI 0.00-0.39), but could increase CT and/or LP testing rates if strictly applied.<sup>39</sup>

## Advanced Imaging

Five studies examined the diagnostic accuracy of non-contrast CT (Figure 4).  $^{20,72-75}$  The pooled estimate for sensitivity was 94% (95% CI 91%-96%,  $I^2 = 74\%$ ) and LR- 0.07 (95% CI 0.03-0.17,  $I^2 = 78\%$ ). van der Wee *et al* assessed CT accuracy at 12 hours, but did not provide sufficient data to reconstruct 2x2 tables stratified by that time threshold. Three other studies reported diagnostic accuracy stratified by the time interval since symptom onset, but only two reported the threshold of less than 6 hours. Pooled sensitivity within the first 6 hours of symptom onset is 100% (95% CI

98%-100%,  $I^2$  = 0%) and LR- 0.01 (95% CI 0-0.04,  $I^2$  = 0%). Beyond 6 hours the pooled sensitivity of CT is 89% (95% CI 83%-93%,  $I^2$  = 89%) and LR- 0.07 (CI 0.01-0.61,  $I^2$  = 63%).

# CSF Erythrocytes and Xanthochromia

Six studies examined diagnostic accuracy of xanthrochromia on CSF analysis, but investigators used variable methods to assess xanthochromia ranging from visible inspection  $^{33,67,68,76}$  to multiple different spectrophotometric assays (Table 5, Figure 5).  $^{34,67,68,77}$  For visible xanthochromia the pooled metanalysis results were: sensitivity 85% (95% CI 66%-96%,  $I^2 = 0\%$ ), specificity 97% (95% CI 96%-98%,  $I^2 = 13\%$ ), LR+ 24.67 (95% CI 12.13-50.14,  $I^2 = 64\%$ ), and LR- 0.22 (95% CI 0.09-0.54,  $I^2 = 13\%$ ). Both Perry and Gangloff evaluated spectrophotometric bilirubin using the United Kingdom National External Quality Assessment Service (UKNEQUAS) algorithm with pooled sensitivity 100% (95% CI 59%-100%,  $I^2 = 0\%$ ), specificity 95% (95% CI 93%-96%,  $I^2 = 98\%$ ), LR+ 15.23 (95% CI 1.58-146.73,  $I^2 = 96\%$ ), and LR- 0.13 (0.02-0.83,  $I^2 = 0\%$ ). The diagnostic accuracies of other spectrophotometric methods to evaluate for CSF bilirubin are also summarized in Table 5.

#### **Test-Treatment Threshold**

The risks of treatment (designated  $R_{rx}$  by Pauker and Kassirer<sup>47</sup>) were primarily those ascribed to angiography itself, and not those of eventual neurosurgical intervention. These risks included additional exposure to ionizing radiation and to contrast, the risks of vascular access and allergic reactions, and the remote risk of stroke or other end-organ damage. The life-time risk of malignancy secondary to exposure to ionizing radiation is a function of dose and age; an adult is typically exposed to 20 mGy radiation during cranial CT. Using an average lifetime cancer risk attributable to such an exposure of almost 1%, 81 we deemed the total risks related to angiography to be relatively low at 0.02.

When considering the benefits of treatment ( $B_{rx}$ ), we considered primarily the benefits of treating an aneurysmal SAH. Overall, SAH has 12% mortality prior to arrival in the ED, and 40% in-hospital mortality at one month despite treatment. In addition, one-third of survivors suffer a severe neurological deficit. See 83 Neurosurgical and interventional radiology treatment options include clipping and coiling the culprit aneurysm when present, with several large randomized trials reporting long-term poor outcomes in up to 35% of treated subjects. The long-term benefits of early diagnosis and treatment in the subset of patients who present awake and neurologically intact are not well characterized but presumably somewhat better, so we assigned an optimistic overall treatment benefit ( $B_{rx}$ ) of 0.8 with immediate diagnosis.

For LP we considered the common morbidity of a post-dural puncture headache, discomfort, and remote risks of infection, and assigned a test risk  $(R_t)$  of 0.01. The testing thresholds from these estimates of potential risk and potential benefit are depicted in Figure 7. The thresholds bracketing indifference between LP versus either no further testing  $(T_{test})$  or angiography  $(T_{treat})$  were very

narrow, suggesting that LP should only be considered for patients with a pre-LP probability between 2% and 7% (when using visible xanthochromia) or between 2% and 4% (when using CSF erythrocyte count). Lower risk patients need no further testing and avoid LP, while higher risk patients benefit from angiography and again avoid LP.

Of note, these rather narrow pre-LP probability bands correspond to the post-CT disease probability, not the original estimate based on history and physical examination (i.e. not the pre-CT estimate). In other words, one should take the pre-CT probability estimate and incorporate the information gained from a negative CT in a Bayesian fashion. The impact of a negative CT on the post-CT (i.e. pre-LP) probability is also shown in Figure 7, depending on the interval between headache onset and imaging. Thus, a pre-LP probability of 2% to 7% represents a pre-CT prior probability of over 70% (based on history and physical) for a CT obtained within 6 hours, and over 20% for a CT obtained beyond 6 hours. The sensitivity analyses showed that these estimates were stable across a range of utilities and were primarily influenced by the diagnostic accuracy of the tests being considered, as illustrated in the Figures. When the risks of testing (LP) were allowed to approach the risks of treatment (angiography), then LP was no longer indicated at any pre-test probability. Table 6 provides an example illustrating how to use Test Indication curves for diagnostic decision-making.

### **Discussion:**

The results of this systematic review demonstrate that neither the presence nor the absence of commonly cited SAH risk factors from history and physical examination accurately rule-in or rule-out this potentially lethal diagnosis. Specifically, no single history or physical exam finding significantly increases (LR+>10) or decreases (LR-<0.1) the post-test probability of SAH for severe headaches that peak within one hour of onset. In addition, many elements of history and physical exam for SAH have only fair to good inter-physician reliability, with the characterization of the headache as "thunderclap" being one of the least reproducible findings. Nonetheless, many physicians would

prefer to never miss a SAH, and missed SAH remains an important cause of litigation in emergency medicine. <sup>23,87</sup>

Lacking accurate features on clinical presentation and recognizing the tension between malpractice risk and a growing imperative to preserve limited healthcare resources, to adhere to Choosing Wisely recommendations, <sup>23,87,88</sup> and to avoid the unintended harms of over-diagnosis, <sup>30</sup> clinical researchers have endeavored to develop more efficient prediction instruments. Yet the development of highly sensitive instruments requires both large datasets and rigorous adherence to methodological standards. To illustrate, a retrospective validation of Rule 1 using medical record review in Northern California reported a LR- of 0.13 (95% CI 0.03-0.61) and reportedly missed 11 cases (20%) of SAH cases despite cranial CT within 6 hours. <sup>89</sup> Another retrospective study of aneurysmal SAH patients demonstrated 100% sensitivity (95% CI 97.6%-100%) using the combination of negative cranial CT and all negative responses on the Ottawa SAH Rule. <sup>90</sup> However, validating a clinical decision rule retrospectively is sub-optimal since all components of the rule may not be recorded in the medical records and whether to code missing variables as unknown or not present is frequently a question. <sup>91</sup> In the retrospective validation of Rule 1, the "onset during exertion" variable was missing in one-third of patients. <sup>92</sup> Therefore, these instruments should ideally undergo prospective validation in different settings before widespread adoption. <sup>91,93</sup>

Unenhanced CT remains the first-line imaging test in ED headache patients with suspected spontaneous SAH. Within 6 hours of headache onset, CT demonstrates sufficient accuracy to rule-in or to rule-out SAH, notwithstanding the heterogeneity of the estimated specificity and LR+ of the two studies identified. Beyond 6 hours the estimated LR- (95% CI 0.01-0.61) and sensitivity (95% CI 83%-93%) become less precise and concerns regarding heterogeneity (I² 89% and 63%, respectively) allow the interpretation that the accuracy of cranial CT to rule-out SAH could range from very good (LR- 0.01) to unhelpful (LR- 0.61) taking the extremes of the 95% confidence interval. Therefore, the timing of CT relative to symptom onset has remained an important qualifier when estimating the

diagnostic accuracy of the CT. Of course, CT can identify other causes of headache or other abnormalities, and similarly can miss SAH due to cervical arteriovenous malformations. Some researchers in the field stipulate that the CT should be interpreted by an experienced neuroradiologist, although recent research indicates that radiologists at non-academic centers may be equally adept.

Despite advances in CT imaging, substantial debate persists surrounding the necessity of LP following a negative head CT, particularly within the first six hours after onset of headache. 94-98

While CSF can provide crucial diagnostic information for alternative etiologies of headache like meningitis, it remains an imperfect test for SAH. Traumatic (blood-contaminated) LPs occur in 1 of 6 cases. 40 Xanthochromia takes many hours to develop, and is variably defined. In addition, post-LP headaches are common 31.99 despite increasing use of smaller gauge, atraumatic spinal needles and reinserting the stylet before needle removal. 32 Other potential complications include low back pain, brainstem herniation, iatrogenic infection, epidural nerve injury, and spinal hemorrhage. 100,101 These factors combined with patients' negative bias against the procedure and the physician time needed help explain emergency physicians' reluctance to pursue LP despite a long-held belief that this procedure was mandatory in the setting of suspected SAH. 102-105 Our analysis provides new insight into the limited utility of performing LP to rule out SAH after a negative CT. We found that, by incorporating the risks and benefits of both testing and continued investigation into the decision analysis, the benefits of LP are unlikely to outweigh the harm across a wide range of reasonable estimates of pre-test disease prevalence and given the limited diagnostic accuracy of the CSF analysis.

Brunell *et al.* noted that LP provided an alternative diagnosis in 3% of suspected SAH cases, but only altered subsequent management in 0.44% of individuals who had an LP.<sup>34</sup> This equates to 227 LPs (1/0.0044) to identify one central nervous system infection requiring antibiotics in suspected SAH patients. In addition, using spectrophotometry, Brunell *et al.* noted five additional cases of SAH

representing 1.1% of LPs and a Number Needed to LP (NN<sub>LP</sub>) of 91 (1/0.011) to identify one additional SAH. Of note, none of these additional cases of "SAH" were aneurysmal, and none underwent subsequent surgery; thus the NN<sub>LP</sub> approaches infinity for detecting SAH that might benefit from intervention. Perry *et al.* reported 17/1546 false negative CTs (with post-headache onset imaging delays ranging from eight hours to eight days) that were identified by LP obtained at the discretion of the attending emergency physician. However, only six of these individuals underwent neurosurgical intervention (ventricular drain, aneurysm coiling, or clipping), which equates to a NN<sub>LP</sub> = 258 (6/1546 = 0.38%, 1/0.0038 = 258) to identify one aneurysm amenable to neurosurgical intervention following a negative cranial CT. Sayer *et al* reported NN<sub>LP</sub> = 250 to identify one additional aneurysmal SAH. Blok *et al* identified one perimesencephalic bleed amongst 760 CT-negative patients with suspected aneurysmal SAH, which they extrapolated to NN<sub>LP</sub> = 15,200 to identify one additional aneurysmal SAH given that 1 in 20 perimesencephalic bleeds are ultimately linked to an aneurysm (760 x 20 = 15,200).  $^{35,106-108}$ 

Another approach to more definitively rule-out SAH after a non-diagnostic, non-contrast CT is to proceed to CT angiography (CTA) without an LP, an approach that 75% of patients favored in one survey  $^{109,110}$  and supporting a role for shared decision-making in evaluation of suspected SAH.  $^{111}$  Some physicians have even advocated for pre-discharge CTA after a non-diagnostic CT and LP. Carstairs *et al* examined CTA in addition to traditional non-contrast CT. Patients that were CT negative and LP negative underwent CTA to more definitively identify either SAH or cerebral aneurysm.  $^{17}$  After evaluating 103 CT-negative/LP-negative patients, they detected two cerebral aneurysms, which equates to a Number Needed to CTA ( $NN_{CTA}$ ) = 52 to detect one cerebral aneurysm that might or might not be the cause of the presenting headache (2/103 = 1.9%, 1/0.019 = 52). Contemplating the  $NN_{CTA}$  is essential because CTA is neither completely safe, nor always clinically meaningful. CTA complication rates range from 0.25%-1.8%.  $^{112,113}$  Carstairs *et al.* interpreted the presence of cerebral aneurysm by CTA to be equivalent to being disease positive for SAH, despite

being unable to establish whether the aneurysm was the etiology of the patient's headache or an incidental finding.

Indeed, intracranial saccular or berry aneurysms are common, occurring in 1%-2% of the population. <sup>114</sup> Linking a non-ruptured aneurysm detected via CTA to an individual patient's headache is problematic. The discovery of an aneurysm in the context of headache is often considered to be a symptomatic aneurysm, and neurosurgical intervention may be advised depending on aneurysm size, location, and specific patient characteristics. <sup>115</sup> However, the relative risk of rupture for cerebral aneurysms detected in symptomatic individuals is 8-fold higher than the rupture rate in asymptomatic individuals with incidentally noted cerebral aneurysms. <sup>116</sup> Based upon one recent meta-analysis of 50 studies, the pooled sensitivity and specificity for CTA to detect aneurysms are 98% and 100%, respectively. <sup>44</sup> MRA can also be utilized to detect intracranial aneurysm, with pooled sensitivity and specificity of 95% and 89% respectively. Although MRA is more time-consuming and generally less available from the ED, it may be useful in circumstances where radiation exposure is to be avoided or a patient has an iodine contrast dye allergy. <sup>45</sup>

Given the risks of over-investigation and incidental findings, it is important to scrutinize the testing thresholds for CT, LP and angiography. We found that only patients with a pre-CT likelihood of SAH well above 20% were likely to benefit from an LP after a negative unenhanced CT. Such pre-CT probabilities are exceptionally high and are far higher than the average diagnostic yield of around 5% for SAH on initial CT. The limited ability of any feature on history or physical examination to identify SAH also renders such high pre-CT probabilities unlikely, except for severe cases likely to have extensive bleeding visible on CT. Moreover, patients with only slightly higher pre-CT probabilities benefit more from angiography than LP, given the limitations of CSF analysis.

Accordingly, this analysis suggests that LP after a negative CT to "rule out" SAH no longer seems rational for the large majority of patients.

The utilitarian approach to decision making test thresholds is considered by some too arbitrary or artificial, yet it is critically important to recognize how to apply these analyses at the bedside. The separation between the test positive and test negative curves are a visual representation of the diagnostic performance of the test, and thus illustrate the summary findings from this meta-analysis. The vertical location of the horizontal line that is used to calculate the test thresholds where it intersects each curve is determined by the ratio between benefits and risks of further treatment, and can easily be shifted up or down to visually estimate new thresholds on the graph. We used a 40:1 ratio of the benefits:risk of angiography, but allowed this to range from 20:1 to 80:1 in a sensitivity analysis and obtained similar findings. We extended this graphical approach to include the non-zero risks of testing, as illustrated by a narrowing of the range of pre-test probabilities for which the test is beneficial. We selected a test:treatment risk of 1:4 for CT and 1:2 for LP compared to angiography, but reduced the risk by half in another sensitivity analysis with again similar findings.

Yet the most compelling argument in support of this analysis is that the findings are strikingly similar to experienced clinicians' own practice. Initial unenhanced CT is rational across a wide range of pretest probabilities, as low as 1%. This 1 in 100 miss rate is consistent with surveys of emergency physicians and a threshold used in clinical decision rule construction for SAH and other serious conditions. <sup>38,104</sup> On the other hand, after a negative CT, the low LR- coupled with a low pre-CT probability for SAH renders the post-CT probability so low that the only moderately accurate LP is rarely warranted. As such, the observed practice of forgoing LP after a negative CT, long considered to be a pitfall in the management of these patients, is easily justified on Bayesian grounds, as represented graphically on the test indication curves.

The test-indication curves also help clinicians avoid the common fallacy sometimes invoked to persuade headache patients that an LP is advisable. Many explain a "95% sensitivity" of CT by

counseling the patient that "there is still a 1 in 20 (i.e. 100% - 95% = 5%) chance that you have a ruptured aneurysm". This is only true if the patient were certain to have a SAH prior to the test being ordered (pre-test probability 100%). It is implausible that even the most skilled clinician can predict SAH prior to any testing in more than 1 in 5 awake and alert patients, the very patient in whom an informed consent discussion is needed prior to LP. Our meta-analysis confirms that there are few findings on history or physical exam with a sufficiently strong association with SAH, and clinical decision rules suffer from modest to poor specificity precisely because of this difficulty identifying disease positive patients prior to testing. Arguably, if one's own diagnostic yield for SAH on initial CT approaches 20% in awake headache patients, one might expect one's miss rate may be unacceptably high. Moreover, such high-risk patients benefit from proceeding directly to CTA especially if the unenhanced CT cannot be done within 6 hours of headache onset, also as indicated on the test threshold curves. On the other hand, using a more reasonable pre-CT probability of 1 in 10 (still higher than the 7% prevalence in prospective studies of emergency patients with headache), a test with 95% sensitivity and perfect specificity implies that the patient being counseled actually has only a 1 in 180 chance of a ruptured aneurysm (post-test odds = pre-test odds x LR- = 1:9 x 0.05).

A cost-effectiveness analysis comparing CT-only, CT/CTA, CT/LP, and CT/MRA as diagnostic modalities found that CT-only and CT/LP are the preferred strategies of diagnosing ED patients with suspected SAH. The CT/LP strategy was slightly more costly, but more effective due to the LP reducing the possibility of a missed diagnosis and subsequent harm. However, since the results of this meta-analysis indicate a much lower threshold for LP than used in the sensitivity analyses of the cost-effectiveness analysis, an updated cost-effectiveness analysis would likely shift the preference towards a CT-only strategy by decreasing the effectiveness of the CT/LP strategy. Further, this cost-effectiveness analysis did not consider crowding in the ED or physician and patient reluctance towards LP, which are practical barriers to the implementation of a CT/LP strategy.

LP has long been considered to be the criterion standard for diagnosing SAH as it may detect small amounts of xanthochromia or blood in the CSF missed by CT. Due to the advancement of multislice CT, sensitivity to detect SAH has increased substantially. The primary clinical question for emergency physicians remains whether LP is indicated to rule out SAH if head CT is negative. Recent literature suggests within 6 hours of symptom onset sensitivity of CT approaches 100%. Additional studies in heterogeneous settings (rural, non-North American, community hospitals) are needed to confirm this diagnostic accuracy.

The significant inter-study heterogeneity observed in our meta-analysis indicates that more research is needed to quantify the diagnostic accuracy of history and physical exam, as well as visible or spectrophotometric assessment of xanthochromia for the diagnosis of SAH. Unfortunately, few SAH studies adhere to the Standards for Reporting of Diagnostic Accuracy (STARD) reporting guidelines. Failure to use reporting standards makes diagnostic studies more difficult for clinicians, researchers, and guideline developers to locate, interpret, and compare. This is reflected in the poor inter-rater reliability results observed in our QUADAS-2 assessments of blinding and appropriately reproducible conduct of the index tests. In addition, studies that do not use the STARD criteria are less likely to report key details necessary for clinicians to understand the risk of bias and expected skew of observed point estimates. Future diagnostic studies should adhere to the STARD reporting standards and history/physical exam studies should follow the STARD criteria recommended for such studies.

Spectrophotometric assessment of xanthochromia should be standardized in terms of conduct and reporting.<sup>78</sup> Most studies did not obtain an LP in all patients or report any standard follow-up protocol

to capture false-negatives. Opening pressures were only obtained in 2% of suspected SAH cases<sup>34</sup> and no studies evaluated the diagnostic accuracy of opening pressures. Future studies reporting the diagnostic accuracy of CT should uniformly report the resolution and timing of imaging, while incorporating post-ED follow-up at reasonable intervals including review of death registries for detection of false-negative results. In addition, the volume of CSF blood and rapidity of aneurysmal bleeding probably influence both CT and LP test accuracy, threatening test independence implicit in Bayesian analyses, and therefore merits assessment for interactions. Future diagnostic researchers should also evaluate test performance stratified by age since aneurysmal SAH are rare in young patients. Furthermore, the skew of observed sensitivities and specificities as a result of spectrum bias, partial verification bias, differential verification bias, and imperfect gold standard bias should be contemplated in future studies.<sup>51</sup>

A planned secondary outcome of this systematic review was to examine clinically significant SAH cases, defined as requiring neurosurgical intervention. The studies in this systematic review did not report adequately to be able to assess these outcomes. Specifically, perimesencephalic hemorrhage is traditionally considered to be "true positive" in SAH diagnostic research, but has a benign clinical course without specific interventions to alter the natural course of disease other than analgesic symptom control. However, perimesencephalic SAH can recur<sup>121</sup> and some neurosurgeons recommend catheter angiography in suspected SAH patients with xanthochromia after a non-diagnostic CTA since a culprit aneurysm will be identified in 8.3% of these cases. <sup>122</sup> In clinical practice SAH is deemed perimesencephalic if two catheter angiography studies 7-days apart demonstrate no cerebral aneurysm, since a single catheter angiogram and CTA can miss 4.2% of aneurysms. <sup>123</sup> None of the diagnostic studies included in this meta-analysis reported uniform CT or catheter angiograms in all patients, let alone repeat catheter angiograms in those with initially "negative" or non-diagnostic advanced imaging. Future studies examining SAH should distinguish aneurysmal SAH from perimesencephalic hemorrhage to better examine SAH that merits surgical intervention.

#### Limitations

Identification of studies for this meta-analysis used only one individual with adjudication of uncertainty of inclusion/exclusion criteria by a second author, which could introduce a bias in study selection. In addition, we excluded non-English studies which could have limited the scope of our findings and yield biased summary estimates, although review of interventional meta-analyses do not suggest significant differences when non-English studies are excluded. 124 Of the 22 studies that provided sufficient detail to be able to reconstruct 2x2 tables, there was significant heterogeneity in the manner in which data were collected and reported. For example, when examining sensitivity and specificity of non-contrast head CT, not all studies reported the specific technology utilized. When technology utilized was reported, it was done so in an inconsistent manner making it unclear which generation (how many slice) CT scanner was used. Variation in technology undoubtedly affects diagnostic test performance, and standardized reporting of the particular technology used should be encouraged in all future studies. 118

Significant variation existed in mechanisms of follow-up for patients that were deemed SAH negative. Some studies did not conduct any follow-up, while others had structured telephone questionnaires, examined coroner records and public death certificates to ensure that patients deemed disease free did not later turn out to have SAH. Without follow-up for patients that were deemed SAH negative it is impossible to know that the patient truly did not have a small CT undetectable SAH or cerebral aneurysm when they were discharged from the ED. Finally, because some of the studies included patients who were unconscious or had neurological deficits, the observed estimates of sensitivity and specificity may be influenced by the spectrum of disease severity.<sup>53</sup>

There was also significant variation in definition of disease, particularly in the incorporation of CSF findings into the definition of being SAH positive. Definition of disease positive ranged from number of red blood cells in the final tube via LP to xanthochromia of CSF. Number of red blood cells was not standardized. Xanthochromia was defined by visual inspection in some studies, and spectrophotometry in others. Xanthochromia via spectrophotometry had substantial variation in the absorption spectrum used to define a positive LP. Other studies used spectrophotometry to detect bilirubin and oxyhemoglobin as indicating a patient with SAH.

### **Conclusions**

This meta-analysis identified a 7.5% weighted prevalence of SAH in ED patients with concerning headache. Although the available high quality diagnostic evidence is limited, there is no single characteristic on history or physical examination that is adequate to rule in or rule out SAH in ED settings. Using a threshold of 1,000 x 10<sup>6</sup>/L RBCs in the final tube of has a summary LR+ 5.7 and LR- 0.21. Within 6 hours of symptom onset, non-contrast cranial CT accurately both rules in and rules out SAH. Given the risks and benefits of further testing and the limited diagnostic accuracy of CSF erythrocyte counts and of various xanthochromia definitions, LP is likely to benefit only patients within a narrow band of pre-LP probabilities, around 2% to 7%. Given the accuracy of CT, even several hours after headache onset, such pre-LP probabilities correspond to pre-CT probabilities of 20% or higher. While many still consider CT followed by LP to be the gold standard for the diagnosis of SAH, few studies use this gold standard given practice patterns evolving away from routine LP. The role of LP may well wane further with improving multi-slice CT and increasing evidence to refute the utility of LP. Validation of SAH clinical decision rules offers the opportunity to more accurately risk-stratify ED headache patients in order to identify subsets most likely to benefit from post-CT LP, CTA, or no further testing.

# References:

- 1. Edlow JA, Panagos PD, Godwin SA, Thomas TL, Decker WW, American College of Emergency P. Clinical policy: critical issues in the evaluation and management of adult patients presenting to the emergency department with acute headache. *Ann Emerg Med.* Oct 2008;52(4):407-436.
- **2.** Landtblom AM, Fridriksson S, Boivie J, Hillman J, Johansson G, Johansson I. Sudden onset headache: a prospective study of features, incidence and causes. *Cephalagia*. Jun 2002;22(5):354-360.
- **3.** Delasobera BE, Osborn SR, Davis JE. Thunderclap headache with orgasm: a case of basilar artery dissection associated with sexual intercourse. *J Emerg Med.* Jul 2012;43(1):e43-e47.
- **4.** Imperato J, Burstein J, Edlow JA. Benign exertional headache. *Ann Emerg Med.* Jan 2003;41(1):98-103.
- **5.** Dodick DW, Wijdicks EF. Pituitary apoplexy presenting as a thunderclap headache. *Neurology.* May 1998;50(5):1510-1511.
- de Bruijn SF, Stam J, Kappelle LJ. Thunderclap headache as first symptom of cerebral venous sinus thrombosis. CVST Study Group. *Lancet* Dec 1996;348(9042):1623-1625.
- 7. Pascual J, Iglesias F, Oterino A, Vasquez-Barquero A, Berciano J. Cough, exertional, and sexual headaches: an analysis of 72 benign and symptomatic cases. *Neurology.* Jun 1996;46(6):1520-1524.
- **8.** Cordenier A, De Hertogh W, De Keyser J, Versijpt J. Headache associated with cough: a review. *J Headache Pain*. May 2013;14:42.
- **9.** Dodick DW, Brown RD, Jr., Britton JW, Huston J, 3rd. Nonaneurysmal thunderclap headache with diffuse, multifocal, segmental, and reversible vasospasm. *Cephalagia*. Mar 1999;19(2):118-123.
- **10.** Day JW, Raskin NH. Thunderclap headache: symptom of unruptured cerebral aneurysm. *Lancet.* Nov 1986;2(8518):1247-1248.
- **11.** Edlow JA. Diagnosis of subarachnoid hemorrhage in the emergency department. *Emerg Med Clin North Am.* Feb 2003;21(1):73-87.
- **12.** Canneti B, Mosqueira AJ, Nombela F, Gilo F, Vivancos J. Spontaneous Subarachnoid Hemorrhage with Negative Angiography Managed in a Stroke Unit: Clinical and Prognostic Characteristics. *J Stroke Cerebrovasc Dis.* Nov 2015;24(11):2484-2490.
- **13.** Elhadi AM, Zabramski JM, Almefty KK, et al. Spontaneous subarachnoid hemorrhage of unknown origin: hospital course and long-term clinical and angiographic follow-up. *J Neurosurg.* Mar 2015;122(3):663-670.

28.

- **14.** Wang QT, Tuhrim S. Etiologies of intracerebral hematomas. *Curr Atherscler Rep.* Aug 2012;14(4):314-321.
- **15.** Hop JW, Rinkel GJ, Algra A, van Gijn J. Case-fatality rates and functional outcome after subarachnoid hemorrhage: a systematic review. *Stroke*. Mar 1997;28(3):660-664.
- **16.** Nieuwkamp DJ, Setz LE, Algra A, Linn FH, de Rooij NK, Rinkel GJ. Changes in case fatality of aneurysmal subarachnoid haemorrhage over time, according to age, sex, and region: a meta-analysis. *Lancet Neurol.* Jul 2009;8(7):635-642.
- **17.** Carstairs SD, Tanen DA, Duncan TD, et al. Computed tomographic angiography for the evaluation of aneurysmal subarachnoid hemorrhage. *Acad Emerg Med.* May 2006;13(5):486-492.
- **18.** Powell J, Kitchen N, Heslin J, Greenwood R. Psychosocial outcomes at 18 months after good neurological recovery from aneurysmal subarachnoid haemorrhage. *J Neurol Neurosurg Psychiatry*. Aug 2004;75(8):1119-1124.
- **19.** Hillman J, Saveland H, Jakobsson KE, et al. Overall management outcome of ruptured posterior fossa aneurysms. *J Neurosurg.* Jul 1996;85(1):33-38.
- **20.** Perry JJ, Stiell IG, Sivilotti ML, et al. Sensitivity of computed tomography performed within six hours of onset of headache for diagnosis of subarachnoid haemorrhage: prospective cohort study. *Bmj.* 2011;343:d4277.
- **21.** Edlow JA, Caplan LR. Avoiding pitfalls in the diagnosis of subarachnoid hemorrhage. *N Engl J Med.* Jan 6 2000;342(1):29-36.
- **22.** Leblanc R. The minor leak preceding subarachnoid hemorrhage. *J Neurosurg.* Jan 1987;66(1):35-39.
- **23.** Johansson A, Lagerstedt K, Asplund K. Mishaps in the management of stroke: a review of 214 complaints to a medical responsibility board. *Cerebrovasc Dis.* 2004;18(1):16-21.
- **24.** Duffy GP. The "warning leak" in spontaneous subarachnoid haemorrhage. *Med J Aust.* May 28 1983;1(11):514-516.
- **25.** Verweij RD, Wijdicks EF, van Gijn J. Warning headache in aneurysmal subarachnoid hemorrhage. A case-control study. *Arch Neurol*. Sep 1988;45(9):1019-1020.
- **26.** Vermeulen MJ, Schull MJ. Missed diagnosis of subarachnoid hemorrhage in the emergency department. *Stroke.* Apr 2007;38(4):1216-1221.
- **27.** Robbert M, Hoogmoed J, van Straaten HA, Coert BA, Peter-Vandertop W, Verbaan D. Time intervals from aneurysmal subarachnoid hemorrhage to treatment and factors contributing to delay. *J Neurol.* Mara 2014;261(3):473-479.
- **28.** Kwiatkowski T, Alagappan K. Headache. In: Marx JA, Hockberger RS, Walls RM, et al., eds. *Rosen's Emergency Medicine: Concepts and Clinical Practice*. 2nd ed: Mosby-Elsevier; 2010:1356-1366.

**30**. **32**. **34**. 39.

- **29.** Hackman JL, Johnson MD, Ma OJ. Spontaneous subarachnoid and intracerebral hemorrhage. In: Tintinalli JE, Stapczynski JS, Ma OJ, Cline DM, Cydulka RK, Meckler GD, eds. *Tintinalli's Emergency Medicine: A Comprehensive Study Guide*. 7th ed. New York City, NY: McGraw-Hill; 2011:1118-1122.
- Carpenter CR, Raja AS, Brown MD. Overtesting and the Downstream Consequences of Overtreatment: Implications of "Preventing Overdiagnosis" for Emergency Medicine *Acad Emerg Med.* Dec 2015 22(12):1484-1492.
- **31.** Seupaul RA, Somerville GG, Viscusi C, Shepard AJ, Hauter WE. Prevalence of postdural puncture headache after ED performed lumbar puncture. *Am J Emerg Med.* Nov 2005;23(7):913-915.
- **32.** Straus SE, Thorpe KE, Holroyd-Leduc J. How do I perform a lumbar puncture and analyze the results to diagnose bacterial meningitis? *Jama*. Oct 25 2006;296(16):2012-2022.
- **33.** Czuczman AD, Thomas LE, Boulanger AB, et al. Interpreting red blood cells in lumbar puncture: distinguishing true subarachnoid hemorrhage from traumatic tap. *Acad Emerg Med.* Mar 2013;20(3):247-256.
- 34. Brunell A, Ridefelt P, Zelano J. Differential diagnostic yield of lumbar puncture in investigation of suspected subarachnoid haemorrhage: a retrospective study. *J Neurol* Jun 2013;260(6):1631-1636.
- **35.** Blok KM, Rinkel GJ, Majoie CB, et al. CT within 6 hours of headache onset to rule out subarachnoid hemorrhage in nonacademic hospitals. *Neurology*. May 2015;84(19):1927-1932.
- **36.** Sayer D, Bloom B, Fernando K, et al. An Observational Study of 2,248 Patients Presenting With Headache, Suggestive of Subarachnoid Hemorrhage, Who Received Lumbar Punctures Following Normal Computed Tomography of the Head. *Acad Emerg Med.* Nov 2015;22(11):1267-1273.
- **37.** Euerle BD. Spinal Puncture and Cerebral Spinal Fluid Examination. In: Roberts JR, Hedges JR, Custalow CB, Chanmugam AS, Chudnofsky CR, MCManus J, eds. *Clinical Procedures in Emergency Medicine*. 5th ed. Philadelphia, PA: Saunders-Elsevier; 2010:1107-1127.
- **38.** Perry JJ, Stiell IG, Sivilotti ML, et al. High risk clinical characteristics for subarachnoid haemorrhage in patients with acute headache: prospective cohort study. *BMJ*. Oct 2010;341:c5204.
- **39.** Perry JJ, Stiell IG, Sivilotti ML, et al. Clinical decision rules to rule out subarachnoid hemorrhage for acute headache. *JAMA*. Sep 25 2013;310(12):1248-1255.
- **40.** Shah KH, Richard KM, Nicholas S, Edlow JA. Incidence of traumatic lumbar puncture. *Acad Emerg Med.* Feb 2003;10(2):151-154.

- **41.** Roost KT, Pimstone NR, Diamond I, Schmid R. The formation of cerebrospinal fluid xanthochromia after subarachnoid hemorrhage. Enzymatic conversion of hemoglobin to bilirubin by the arachnoid and choroid plexus. *Neurology*. Sep 1972;22(9):973-977.
- **42.** Edlow JA, Bruner KS, Horowitz GL. Xanthochromia. *Arch Pathol Lab Med.* Apr 2002;126(4):413-415.
- 43. Chu K, Hann A, Greenslade J, Williams J, Brown A. Spectrophotometry or Visual Inspection to Most Reliably Detect Xanthochromia in Subarachnoid Hemorrhage: Systematic Review. *Ann Emerg Med.* Sep 2014;64(3):256-264.e255.
- **44.** Westerlaan HE, van Dijk JM, Jansen-van der Weide MC, et al. Intracranial aneurysms in patients with subarachnoid hemorrhage: CT angiography as a primary examination tool for diagnosis--systematic review and meta-analysis. *Radiology.* Jan 2011;258(1):134-145.
- **45.** Sailer AM, Wagemans BA, Nelemans PJ, de Graaf R, van Zwam WH. Diagnosing intracranial aneurysms with MR angiography: systematic review and meta-analysis. *Stroke.* Jan 2014;45(1):119-126.
- **46.** Pauker SG, Kassirer JP. Therapeutic decision making: a cost-benefit analysis. *N Engl J Med.* Jul 1975;293(5):229-234.
- **47.** Pauker SG, Kassirer JP. The threshold approach to clinical decision making. *N Engl J Med.* May 15 1980;302(20):1109-1117.
- **48.** Bernstein J. Test-indication curves. *Med Decis Making* Jan-Mar 1997;17(1):103-106.
- **49.** Liberati A, Altman DG, Tetzlaff J, et al. The PRISMA statement for reporting systematic reviews and meta-analyses of studies that evaluate healthcare interventions: explanation and elaboration. *Bmj.* 2009;339:b2700.
- **50.** Stroup DF, Berlin JA, Morton SC, et al. Meta-analysis of observational studies in epidemiology: a proposal for reporting. Meta-analysis Of Observational Studies in Epidemiology (MOOSE) group. *Jama*. Apr 19 2000;283(15):2008-2012.
- **51.** Kohn MA, Carpenter CR, T.B. N. Understanding the direction of bias in studies of diagnostic test accuracy. *Acad Emerg Med.* Nov 2013;20(11):1194-1206.
- **52.** Whiting PF, Rutjes AW, Westwood ME, et al. QUADAS-2: a revised tool for the quality assessment of diagnostic accuracy studies. *Ann Intern Med.* Oct 18 2011;155(8):529-536.
- **53.** Mulherin SA, Miller WC. Spectrum bias or spectrum effect? Subgroup variation in diagnostic test evaluation. *Ann Intern Med.* Oct 2002;137(7):598-602.
- **54.** Tsementzis SA, Hitchcock ER, DeCothi A, Gill JS. Comparative studies of the diagnostic value of cerebrospinal fluid spectrophotometry and computed tomographic scanning in subarachnoid hemorrhage. *Neurosurgery*. Dec 1985;17(6):908-912.

57. **58. 59**. 60. 61. **62**. 63. 66. 67. 68.

- **55.** Kanzaria HK, McCabe AM, Meisel ZM, et al. Advancing Patient-centered Outcomes in Emergency Diagnostic Imaging: A Research Agenda. *Acad Emerg Med.* Dec 2015;22(12):1435-1446.
- Byrt T, Bishop J, Carlin JB. Bias, prevalence and kappa. *J Clin Epidemiol.* May 1993;46(5):423-429.
- **57.** Macaskill P GC, Deeks JJ, Harbord RM,, Y T. *Cochrane Handbook for Systematic Reviews of Diagnostic Test Accuracy Version 0.9.0.* The Cochrane Collaboration,; 2010.
- **58.** Zamora J, Abraira V, Muriel A, Khan K, Coomarasamy A. Meta-DiSc: a software for meta-analysis of test accuracy data. *BMC Med Res Methodol.* 2006;6:31.
- **59.** Der Simonian R, Laird N. Meta-analysis in clinical trials. *Control Clin Trials.* Sep 1986;7(3):177-188.
- **60.** Der Simonian R, Kacker R. Random-effects model for meta-analysis of clinical trials: an update. *Contemp Clin Trials*. Feb 2007;28(2):105-114.
- **61.** Deeks JJ, Macaskill P, Irwig L. The performance of tests of publication bias and other sample size effects in systematic reviews of diagnostic test accuracy was assessed. *J Clin Epidemiol.* Sep 2005;58(9):882-893.
- **62.** Grimes DA, Schulz KF. Refining clinical diagnosis with likelihood ratios. *Lancet.* Apr 2005;365(9469):1500-1505.
- **63.** Gallagher EJ. Clinical utility of likelihood ratios. *Ann Emerg Med.* Mar 1998;31(3):391-397.
- **64.** Hayden SR, Brown MD. Likelihood ratio: A powerful tool for incorporating the results of a diagnostic test into clinical decisionmaking. *Ann Emerg Med.* May 1999;33(5):575-580.
- **65.** Glasziou P. Threshold analysis via the Bayes' nomogram. *Med Decis Making.* 1991;11(1):61-62.
- **66.** Byrt T. How good is that agreement? *Epidemiology.* 1996;7(5):561.
- **67.** Perry JJ, Sivilotti ML, Stiell IG, et al. Should spectrophotometry be used to identify xanthochromia in the cerebrospinal fluid of alert patients suspected of having subarachnoid hemorrhage? *Stroke*. Oct 2006;37(10):2467-2472.
- **68.** Gangloff A, Nadeau L, Perry JJ, Baril P, Emond M. Ruptured aneurysmal subarachnoid hemorrhage in the emergency department: Clinical outcome of patients having a lumbar puncture for red blood cell count, visual and spectrophotometric xanthochromia after a negative computed tomography. *Clin Biochem.* Jul 2015;48(10-11):634-639.
- **69.** Linn FH, Rinkel GJ, Algra A, van Gijn J. Headache characteristics in subarachnoid haemorrhage and benign thunderclap headache. *J Neurol Neurosurg Psychiatry*. Nov 1998;65(5):791-793.

- **70.** Morgenstern LB, Luna-Gonzales H, Huber JC, et al. Worst headache and subarachnoid hemorrhage: prospective, modern computed tomography and spinal fluid analysis. *Ann Emerg Med.* Sep 1998;32(3 Pt 1):297-304.
- **71.** Bø SH, Davidsen EM, Gulbrandsen P, Dietrichs E. Acute headache: a prospective diagnostic work-up of patients admitted to a general hospital. *Eur J Neurol.* Dec 2008;15(12):1293-1299.
- **72.** van der Wee N, Rinkel GJ, Hasan D, van Gijn J. Detection of subarachnoid haemorrhage on early CT: is lumbar puncture still needed after a negative scan? *J Neurol Neurosurg Psychiatry*. Mar 1995;58(3):357-359.
- **73.** Boesiger BM, Shiber JR. Subarachnoid hemorrhage diagnosis by computed tomography and lumbar puncture: are fifth generation CT scanners better at identifying subarachnoid hemorrhage? *J Emerg Med.* Jul 2005;29(1):23-27.
- **74.** O'Neill J, McLaggan S, Gibson R. Acute headache and subarachnoid haemorrhage: a retrospective review of CT and lumbar puncture findings. *Scott Med J.* Nov 2005;50(4):151-153.
- **75.** Backes D, Rinkel GJ, Kemperman H, Linn FH, Vergouwen MD. Time-dependent test characteristics of head computed tomography in patients suspected of nontraumatic subarachnoid hemorrhage. *Stroke.* Aug 2012;43(8):2115-2119.
- **76.** Dupont SA, Wijdicks EF, Manno EM, Rabinstein AA. Thunderclap headache and normal computed tomographic results: value of cerebrospinal fluid analysis. *Mayo Clinic Proc.* Dec 2008;83(12):1326-1331.
- 77. Hann A, Chu K, Greenslade J, Williams J, Brown A. Benefit of cerebrospinal fluid spectrophotometry in the assessment of CT scan negative suspected subarachnoid haemorrhage: a diagnostic accuracy study. *J Clin Neurosci.* Jan 2015;22(1):173-179.
- **78.** Cruickshank A, Auld P, Beetham R, et al. Revised national guidelines for analysis of cerebrospinal fluid for bilirubin in suspected subarachnoid haemorrhage. *Ann Clin Biochem.* May 2008;45(Pt3):238-244.
- **79.** Perry JJ, Alyahya B, Sivilotti ML, et al. Differentiation between traumatic tap and aneurysmal subarachnoid hemorrhage: prospective cohort study. *BMJ*. Feb 2015;350:h568.
- **80.** Brown MD, Reeves MJ. Evidence-based emergency medicine skills for evidence-based emergency care. Interval likelihood ratios: another advantage for the evidence-based diagnostician. *Ann Emerg Med.* Aug 2003;42(2):292-297.
- **81.** Hall EJ, Brenner DJ. Cancer risks from diagnostic radiology. *Br J Radiol.* May 2008;81(965):362-378.
- **82.** Schievink WI. Intracranial aneurysms. *N Engl J Med.* Jan 2 1997;336(1):28-40.
- **83.** Burrows AM, Korumilli R, Lanzino G. How we do it: acute management of subarachnoid hemorrhage. *Neurol Res.* Mar 2013;35(2):111-116.

86. 92. 93. 95. 96.

- **84.** McDougall CG, Spetzler RF, Zabramski JM, et al. The Barrow Ruptured Aneurysm Trial. *J Neurosurg.* Jan 2012;116(1):135-144.
- **85.** Spetzler RF, McDougall CG, Albuquerque FC, et al. The Barrow Ruptured Aneurysm Trial: 3-year results. *J Neurosurg.* Jul 2013;119(1):146-157.
- Molyneux AJ, Kerr RS, Yu LM, et al. International subarachnoid aneurysm trial (ISAT) of neurosurgical clipping versus endovascular coiling in 2143 patients with ruptured intracranial aneurysms: a randomised comparison of effects on survival, dependency, seizures, rebleeding, subgroups, and aneurysm occlusion. *Lancet.* Sep 2005;366(9488):809-817.
- **87.** McNeill A. Neurological negligence claims in the NHS from 1995 to 2005. *Eur J Neurol.* Apr 2007;14(4):399-402.
- **88.** Choosing Wisely: American College of Emergency Physicians. 2014; http://www.choosingwisely.org/societies/american-college-of-emergency-physicians/. Accessed October 16, 2015.
- **89.** Mark DG, Hung YY, Offerman SR, et al. Nontraumatic subarachnoid hemorrhage in the setting of negative cranial computed tomography results: external validation of a clinical and imaging prediction rule. *Ann Emerg Med.* Jul 2013;62(1):1-10.
- **90.** Mark DG, Kene MV, Udaltsova N, Vinson DR, Ballard DW. Sensitivity of a Clinical Decision Rule and Early Computed Tomography in Aneurysmal Subarachnoid Hemorrhage. *Western J Emerg Med.* Sep 2015;16(5):671-676.
- **91.** Hunter BR, Carpenter CR. The development of clinical prediction rules. In: Wilson MP, Guluma KZ, Hayden SR, eds. *Doing Research in Emergency & Acute Care: Making Order Out of Chaos*. Oxford, UK: Wiley-Blackwell; 2015:139-147.
- **92.** Perry JJ, Sivilotti ML, Stiell IG. Pragmatic interpretation of the study of nontraumatic subarachnoid hemorrhage in the setting of negative cranial computed tomography results: external validation of a clinical and imaging prediction rule. *Ann Emerg Med.* Oct 2013;62(4):435-436.
- **93.** Perry JJ, Marcantonio A, Sivilotti MLA. Prospective Validation of the Ottawa SAH Clinical Decision Rule to Exclude Subarachnoid Hemorrhage in Patients with Acute Headache *NEJM*. 2016 (in peer review).
- **94.** Coats TJ, Loffhagen R. Diagnosis of subarachnoid haemorrhage following a negative computed tomography for acute headache: a Bayesian analysis. *Eur J Emerg Med.* Apr 2006;13(2):80-83.
- **95.** Coats TJ. Subarachnoid haemorrhage: lumbar puncture for every negative scan? *BMJ.* Aug 2006;333(7564):396-397.
- **96.** Gee C, Dawson M, Bledsoe J, et al. Sensitivity of newer-generation computed tomography scanners for subarachnoid hemorrhage: a Bayesian analysis. *J Emerg Med.* 2012;43(1):13-18.

- **97.** Malabarey MA, Barbic D. Can the combination of a negative computed tomography result and a negative lumbar puncture safely exclude the diagnosis of subarachnoid hemorrhage in patients with thunderclap headache? *CJEM*. 2013;15(2):113-115.
- **98.** Farzad A, Radin B, Oh JS, et al. Emergency diagnosis of subarachnoid hemorrhage: an evidence-based debate. *J Emerg Med.* May 2013;44(5):1045-1053.
- **99.** Kuntz KM, Kokmen E, Stevens JC, Miller P, Offord KP, Ho MM. Post-lumbar puncture headaches: experience in 501 consecutive procedures. *Neurology*. Oct 1992;42(10):1884-1887.
- **100.** Abouleish E, Vega S, Blendinger I, Tio TO. Long-term follow-up of epidural blood patch. *Anesth Analg.* Jul-Aug 1975;54(4):459-463.
- **101.** Evans RW. Complications of lumbar puncture. *Neurol Clin.* Feb 1998;16(1):83-105.
- **102.** Muhammed O, Teubner D, Jones DN, Slavotinek JP. Retrospective audit of the investigation of patients with suspected acute subarachnoid haemorrhage. *J Med Imaging Radiat Oncol.* Aug 2010;54(4):339-346.
- **103.** Mehrotra P, Sookhoo S, Kolla S, Halbert H, Lavell K, England S. Investigation of subarachnoid haemorrhage: does the buck stop with CT? *J Med Life.* Jul-Sep 2010;3(3):338-342.
- **104.** Perry JJ, Stiell IG, Wells GA, et al. Attitudes and judgment of emergency physicians in the management of patients with acute headache. *Acad Emerg Med.* Jan 2005;12(1):33-37.
- **105.** Perry JJ, Stiell IG, Wells GA, Spacek A. Diagnostic test utilization in the emergency department for alert headache patients with possible subarachnoid hemorrhage. *CJEM*. Sep 2002;4(5):333-337.
- **106.** Rinkel GJ, Wijdicks EF, vermeulen MJ, et al. Nonaneurysmal perimesencephalic subarachnoid hemorrhage: CT and MR patterns that differ from aneurysmal rupture. *AJNR Am J Neuroradiol.* Sep-Oct 1991;12(5):829-834.
- **107.** Topcuoglu MA, Oqilvy CS, B.S. C, Buonanno FS, Koroshetz WJ, Singhal AB. Subarachnoid hemorrhage without evident cause on initial angiography studies: diagnostic yield of subsequent angiography and other neuroimaging tests. *J Neurosurg.* Jun 2003;98(6):1235-1240.
- **108.** Kalra VB, Wu X, Matouk CC, Malhotra A. Use of follow-up imaging in isolated perimesencephalic subarachnoid hemorrhage: a meta-analysis. *Stroke.* Feb 2015;46(2):401-406.
- **109.** Youssef NA, Gordon AJ, Moon TH, et al. Emergency department patient knowledge, opinions, and risk tolerance regarding computed tomography scan radiation. *J Emerg Med.* Feb 2014;46(2):208-214.
- **110.** McCormack RF, Hutson A. Can computed tomography angiography of the brain replace lumbar puncture in the evaluation of acute-onset headache after a negative noncontrast cranial computed tomography scan? *Acad Emerg Med.* Apr 2010;17(4):444-451.

- **111.** Hess EP, Grudzen CR, Thomson R, Raja AS, Carpenter CR. Shared Decision-making in the Emergency Department: Respecting Patient Autonomy When Seconds Count. *Acad Emerg Med.* Jul 2015;22(7):856-864.
- **112.** Linn FH, Wijdicks EF, vander Graaf Y, Weerdesteyn-van Vliet FA, Bartelds AI, van Gijn J. Prospective study of sentinel headache in aneurysmal subarachnoid haemorrhage. *Lancet.* Aug 1994;344(8922):590-593.
- **113.** Cloft HJ, Joseph GJ, Dion JE. Risk of cerebral angiography in patients with subarachnoid hemorrhage, cerebral aneurysm, and arteriovenous malformation: a meta-analysis. *Stroke.* Feb 1999;30(2):317-320.
- **114.** Brown RD, Jr., Broderick JP. Unruptured intracranial aneurysms: epidemiology, natural history, management options, and familial screening. *Lancet Neurol.* Apr 2014;13(4):393-404.
- **115.** Brennan JW, Schwartz ML. Unruptured intracranial aneurysms: appraisal of the literature and suggested recommendations for surgery, using evidence-based medicine criteria. *Neurosurgery*. Dec 2000;47(6):1359-1371; discussion 1371-1352.
- **116.** Rinkel GJ, Djibuti M, Algra A, van Gijn J. Prevalence and risk of rupture of intracranial aneurysms: a systematic review. *Stroke.* Jan 1998;29(1):251-256.
- **117.** Ward MJ, Bonomo JB, Adeoye O, Raja AS, Pines JM. Cost-effectiveness of diagnostic strategies for evaluation of suspected subarachnoid hemorrhage in the emergency department. *Acad Emerg Med.* Oct 2012;19(10):1134-1144.
- **118.** Bossuyt PM, Reitsma JB, Bruns DE, et al. The STARD statement for reporting studies of diagnostic accuracy: explanation and elaboration. *Ann Intern Med.* Jan 2003;138(1):W1-W12.
- **119.** Smidt N, Rutjes AW, van der Windt DA, et al. The quality of diagnostic accuracy studies since the STARD statement: has it improved? *Neurology*. Sep 2006;67(5):792-797.
- **120.** Simel DL, Rennie D, Bossuyt PM. The STARD statement for reporting diagnostic accuracy studies: application to the history and physical examination. *J Gen Intern Med.* Jun 2008;23(6):768-774.
- **121.** Reynolds MR, Blackburn SL, Zipfel GJ. Recurrent idiopathic perimesencephalic subarachnoid hemorrhage. *J Neurosurg.* Sep 2011;115(3):612-616.
- **122.** Wallace AN, Dines JN, Zipfel GJ, Derdeyn CP. Yield of catheter angiography after computed tomography negative, lumbar puncture positive subarachnoid hemorrhage. *Stroke.* Jun 2013;44(6):1729-1731.
- **123.** Delgado-Almandoz JE, Jagadeesan BD, Refai D, et al. Diagnostic yield of repeat catheter angiography in patients with catheter and computed tomography angiography negative subarachnoid hemorrhage. *Neurosurgery*. May 2012;70(5):1135-1142.

**124.** Morrison A, Polisena J, Husereau D, et al. The effect of English-language restriction on systematic review-based meta-analyses: a systematic review of empirical studies. *Int J Technol Assess Health Care.* Apr 2012;28(2):138-144.

# Table 1

Potential Etiologies of Spontaneous SAH

Cerebral aneurysm

Perimesencephalic

Isolated convexity

Vascular malformations (arteriovenous malformations/AVF)

Arterial dissection

Cerebral amyloid angiopathy

Moyamoya

Vasculitis (PRES, RCVS, lupus)

Coagulopathy (thrombocytopenia, anticoagulation)

Sickle cell disease

Hypertension

Sympathomimetic drugs

# **Table 2 Summary of Included Studies**

S	Study	Location	Setting	No. Patients	Mean age <sup>1</sup>	Inclusion Criteria	Study Design	SAH Criterion Standard	SAH prevalence %	Variables Assessed	Follow-Up
Back	kes 2012	University Medical Center, Utrecht, Netherlands, 2005-2012	ED	250	48*	Age >16 years, clinical suspicion of acute onset SAH, GCS 15	Two prospective databases. First database consecutive patients with confirmed SAH. Second all LP patients with SPT	Not explicitly defined	42.0	Sen/Spec CT at <6- hours and >6-hours for SAH	None
Во	o 2008	Akershus University Hospital, Norway, 1998- 2002	ED	433	44	Age >14 years, acute onset worst headache of life	Prospective observational consecutive sampling cohort	Consensus between two neurologists, but not explicitly defined	16.4	Sen/Spec of elements of history including activity at onset, and peak intensity for SAH	None

<sup>&</sup>lt;sup>1</sup> Median

Boesiger 2005	Pitt County Memorial Hospital, Greenville SC, 2002	ED	177	NR	Adults with headache work-up including CT and LP	Retrospective chart review	SAH observed on CT or ≥400 RBC in CSF tube 1 that did not clear by 10-fold	3.4	Sen/Spec of CT for SAH	Phone at 1-ye if ≥40 RBC CSF to 1 that o not cle by 10-f
Brunell 2013	Uppsala University Hospital, Uppsala Sweden, 2009-2011	Hospit al	453	NR	Age>10 years with LP to exclude SAH	Lab information system review	CSF hemoglobin detection via SPT using 415 nm wavelength or CSF bilirubin detection via automatic analyzer	1.1	Sen/spec of CSF RBC for SAH	Non
Carstairs 2006	Naval Medical Center San Diego CA, July 2002-July 2004	ED	116	39	Age>18 with either worst headache of life or onset <60 seconds or associated neuro deficit	Prospective observational consecutive sampling cohort	SAH observed on CT or RBC in CSF tube 1 that did not clear by >25% in tube 4 or presence of xanthochromia on visual	4.3	Sen/Spec of CTA & elements of history including stiff neck, blurred vision, &	Non

rticle								inspection by lab		altered mental status
ted A	Czuczman 2013	Unspecified urban tertiary care hospital, 2001-2009	ED	220	44	Adults with headaches billed for LPs, ≥5 RBC in final CSF tube, and either CT angiogram or magnetic resonance angiogram within 2-weeks	Retrospective case series	Either 1) presence of SAH on imaging; 2) xanthochromia with aneurysm or AVM>2mm;  3) xanthochromia and culture- or PCR-negative meningitis	11.8	Sen/Spec of elements of history including sudden onset headache, worst headache, & onset with exertion; CSF RBC iLRs
C 6 10	Dupont 2008	St. Mary's Hospital, Rochester MN, 1998- 2007	ED	117	45	Age>18 with headache onset <2 weeks prior with both non-contrast CT for detection of blood and LP	Retrospective chart review	Not specified	12.0	Sen/Spec visual xanthochr omia for aneurysm

None

None

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	Gangloff	Hôpital de	ED	706	41	Age>14 with	Retrospective	Presence of	0.7	Sen/Spec	None
	2015	l'Enfant-Jésus,				acute	case series	any		visual	
		Quebec				headache		aneurysm		xantho,	
		Canada, 2003-				suspicious		and presence		iterative	
		2009				for SAH, GCS 15, and		of either		SPT, or	
						initial head		visual xantho		UK	
						CT negative		or >5 x $10^6$		NEQUA	
						for SAH		RBC/L in last		SPT	
						with		CSF tube			
						subsequent					
•						LP					

Hann 2015	Royal Brisbane & Women's Hospital, June 2005-July 2012	ED	409	38	Adults discharged with headache after CSF evaluation for xantho	Retrospective case series	Presence of vascular aneurysm on angiogram within 30- days of headache or no repeat ED visit or SAH death in 30- days	1.5	Sen/Spec visual xantho via visual inspection	Review of Queensland death registry
Landtblom 2002	University Hospital, Linköping Sweden over unspecified	ED	137	42*	Age>18 years presenting within 10 days of headache	Prospective observational cohort	Not specified	11.3	Sen/Spec of elements of history including	Phone at 2-4 days, 7-9 days, and 1-, 3-, 6-, and

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	31-month period				onset with abrupt onset (<10 second) headache, examined by on-call Neurologist				sudden onset headache, worst headache, & nausea	12- months for 90 subjects without SAH
Linn 1998	Utrecht University Hospital, Utrecht, Netherlands, Jan 1992-Oct 1994	ED	102	48	Alert adult patients without focal neuro deficit referred to ED by general practitioner with concern for SAH with CT and LP	Prospective observational cohort	Not specified	41.2	Sen/Spec of elements of history including activity at onset and acuity of onset	None

Morgenstern 1998	Hermann Hospital, Houston TX, March 1995- June 1996	ED	97	39	Adult patients with 10/10 worst headache of life	Prospective observational cohort	Either 1) presence of SAH on imaging OR 2) CSF RBC > 1000 with <25% decrement from first to last tube with either visual xantho, SPT xantho, or elevated D- dimer	18.6	Sen/Sp of elemen of histo & physic exam includir nausea stiff neo and letharg
O'Neill 2005	Unspecified hospital over unspecified 1- year period	ED	116	NR	Acute headache <24 hours, clinical suspicion SAH, referred for CT from ED	Retrospective chart review	Not specified	16.4	Sen/Sp of CT: SAH
Perry 2006	6 Canadian tertiary care EDs, July 2002-Jan 2004	ED	220	42	Age>15 years with acute, spontaneous headache with or without	Pre-planned substudy of prospective, consecutive patient cohort study	Any one of the following: subarachnoid blood on CT, >5x10 <sup>6</sup> /L RBC in 3 <sup>rd</sup> or 4 <sup>th</sup> tube of CSF,	0.9	Sen/Sp of visu inspect for xan and 4 S metho

Phone f/u

at mean 24 months

None

Phone f/u

at 1month for

CT-, LP-

patients;

chart

						syncope		or visible xantho + aneurysm on cerebral angiography		for xantho for SAH
	Perry 2010	6 Canadian tertiary care EDs, Nov 2000-Nov 2005	ED	1999	43	Age≥16 years with spontaneous headache peaking within 1- hour or syncope associated	Prospective cohort study	Any one of the following: subarachnoid blood on CT, visual xantho, >5x10 <sup>6</sup> /L	6.5	Sen/Spec of elements of history including activity at onset, neck
Dte						with a headache		RBC in the final tube of CSF with an aneurysm or AVM on cerebral angiography		stiffness, and arrival mode. Also, derivation of 3 CDRs

					associated with a headache		>5x10 <sup>6</sup> /L RBC in the final tube of CSF with an aneurysm or AVM on cerebral angiography		neck stiffness, and arrival mode. Also, derivation of 3 CDRs	chart review & Ontario registry review at study end for same patients
Perry 2011	11 Canadian tertiary care EDs, Nov 2000-Dec	ED	3132	45	Age>15 years with acute, spontaneous headache	Prospective cohort study	Any one of the following: subarachnoid blood on CT,	7.7	Sen/Spec CT at <6- hours and >6-hours	Phone f/u at 1- and 6-months for all patients
					headache				[	

review at 1-month for all patients

Phone f/u at 1- and 6-months for all patients without CT or LP;

	2009	with or	visual	for SAH	without
		without	xantho,		CT or LP;
		syncope	$>5x10^6/L$		chart
-			RBC in the		review &
			final tube of		provincial
			CSF with an		coroner's
			aneurysm or		records at
			AVM on		study end
			cerebral		for same
			angiography		patients

Perry 2013	10 Canadian tertiary care EDs, Nov 2000-Dec 2009	ED	2131	44	Age≥16 years with spontaneous headache peaking within 1- hour or syncope associated with a headache	Prospective cohort study	Any one of the following: subarachnoid blood on CT, visual xantho, >5x10 <sup>6</sup> /L RBC in the final tube of CSF with an aneurysm or AVM on cerebral angiography	6.2	Sen/Spec of elements of history & physical exam including activity at onset, neck stiffness, and arrival mode. Also, derivation of 4 SAH CDRs	Phone f/u at 1- and 6-months for all patients without CT or LP; chart review & provincial coroner's records at study end for same patients
Perry 2015	12 Canadian tertiary care EDs, April 2006-July 2010	ED	1739 1098 normal LP 641 abnl	44	Age>15 years with acute, spontaneous headache with or without syncope with LP to evaluate potential	Pre-planned substudy of prospective, consecutive patient cohort study	Aneurysmal SAH if: subarachnoid blood on CT, visual xantho,or any RBC in the final tube of CSF with an aneurysm on cerebral	2.3	Sen/Spec of <2000 x 10 <sup>6</sup> /L RBC in CSF + no xantho for diagnosis aneurysmal SAH	Phone f/u at 1- and 6-months for all patients without CT or LP; chart review & provincial coroner's

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	LP	SAH	angiography	records at
				study end
				for same
				patients
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Van der Wee 1995	University hospitals in Utrecht and Rotterdam Netherlands, 1989-1993	NR	175	NR	Acute headache without confusion or focal deficit with CT within 12- hours	Unclear whether retrospective or prospective "consecutive" case series	Not specified	68	Sen/Spec CT at <12-hours for SAH	None
Wood 2015	Princess Alexandra Hospital, Brisbane Queensland Australia, Jan 2000-April 2003	NR	240	NR	Acute or unusually severe headache with LP after a normal cranial CT for possible diagnosis of SAH	Retrospective chart review	Uniform CSF bloodstaining across serial samples with visual xantho and "positive" angiography	0.8	Sen/Spec of CSF SPT (XI > 0.080)	None

**Table 3 Overview of Quality Assessment of Included Studies** 

Study	Sample	Avoided Inappropriat e Exclusions?	Patients and Settings Match Study Question ?	Blinded Interpretatio n of Index Test?	Acceptable Index Test Application ?	Acceptabl e Reference (Gold) Standard?	Outcomes Assessed by Blinded Assessor ?	Assessed Outcome s Pertinent ?	Appropriat e Interval Between Index Test & Reference Standard?	Same Referenc e (Gold) Standard for All Patients?	All Patients Analyzed ?
Backes 2012	Consecutiv e	YES	YES	NO	YES	NO	NO	NO	YES	YES	YES
Bo 2008	Consecutiv e	YES	YES	NO	NO	NO	NO	NO	YES	YES	YES
Boesiger 2005	Consecutiv e	YES	YES	NO	YES	YES	YES	NO	YES	YES	YES
Brunell 2013	Consecutiv e	YES	YES	YES	NO	NO	YES	NO	YES	YES	YES
Carstairs 2006	Consecutiv e	YES	YES	YES	YES	YES	YES	YES	YES	YES	YES
Czuczman 2013	Consecutiv e	YES	NO	YES	YES	YES	NO	YES	YES	YES	YES
Dupont 2008	Consecutiv e	NO	YES	NO	YES	YES	NO	YES	YES	NO	YES
Gangloff 2015	Consecutiv e	YES	YES	NO	NO	YES	NO	YES	YES	NO	NO
Hann 2014	Consecutiv e	NO	NO	NO	NO	YES	NO	YES	YES	YES	YES

Landtblom 2002	Consecutiv e	YES	YES	NO	NO	YES	NO	YES
Linn 1998	Consecutiv e	YES	YES	NO	YES	NO	NO	YES
Morgenster n 1998	Consecutiv e	YES	YES	YES	NO	NO	YES	NO
D'Neill 2005	Consecutiv e	YES	YES	YES	YES	NO	YES	YES
Perry 2006	Consecutiv e	NO	YES	YES	YES	YES	NO	YES
Perry 2010	Consecutiv e	YES	YES	YES	YES	YES	NO	YES
Perry 2011	Consecutiv e	YES						
Perry 2013	Consecutiv e	YES						
Perry 2015	Consecutiv e	YES	YES	YES	YES	YES	NO	YES
van der Wee 1995	Consecutiv e	NO						
Wood 2005	Consecutiv e	NO	NO	NO	NO	NO	NO	YES
Kappa <sup>2</sup>	1.00	0.33	0.52	0.05	0.14	0.05	0.05	0.43
,								

YES

NO

YES

YES

YES

YES

YES

YES

YES

YES

YES

0.62

YES

YES

NO

NO

YES

NO

NO

YES

YES

YES

YES

0.33

<sup>&</sup>lt;sup>2</sup> Prevalence adjusted, bias adjusted kappa as described by .

Risk Factor	Sensitivity, % (95% CI)	Specificity, % (95% CI)	Positive LR (95% CI)	Negative LR (95% CI)
Diplopia				
Landtblom 2002	0 (0-15)	98 (94-100)	0.96 (0.05-19.33)	1.00 (0.94-1.07)
Family History cerebral				
aneurysm				
Czuczman 2013	0 (0-13)	92 (87-95)	0.22 (0.01-3.54)	1.07 (1.00-1.15)
Lethargy				
Morgenstern 1998	39 (17-64)	82 (72-90)	2.19 (1.04-4.64)	0.74 (0.51-1.09)
Onset < 1 minute	50 (34-66)	45 (32-58)	0.91 (0.62-1.33)	1.11 (0.74-1.68)
Linn 1998				
Onset 1-5 minutes	24 (12-39)	87 (75-94)	1.79 (0.77-4.14)	0.88 (0.72-1.07)
Linn 1998				
Onset <1 hour	100 (95-100)	12 (9-16)	1.13 (1.09-1.19)	0.06 (0-0.95)
Bo 2008				
PMH chronic headache	19 (7-39)	79 (73-85)	0.93 (0.40-2.91)	1.02 (0.83-1.24)
Czuczman 2013				
PMH of hypertension	31 (14-52)	80 (74-85)	1.53 (0.81-2.91)	0.87 (0.66-1.13)
Czuczman 2013				
Scotomata	0 (0-15)	93 (87-97)	0.28 (0.02-4.72)	1.06 (0.98-1.14)
Landtblom 2002				
Similar headache in past	19 (9-34)	90 (79-96)	1.90 (0.71-5.09)	0.90 (0.76-1.07)
Linn 1998				

			nd Xanthochromia for S	
Risk Factor	Sensitivity, % (95%	Specificity, % (95%	Positive LR	Negative LR (959
	CI)	CI)	(95% CI)	CI)
CSF RBC >1000				
Czuczman 2013	65 (44-83)	79 (72-84)	3.09 (2.09-4.57)	0.44 (0.26-0.75)
Perry 2015	93 (68-100)	91 (88-93)	10.25 (7.73-13.59)	0.07 (0.01-0.49)
Pooled Accuracy	76 (60-88)	88 (86-90)	5.66 (1.38-23.27)	0.21 (0.03-1.66)
(anthochromia – UK				
NEQAS	100 (16-100)	83 (77-88)	4.87 (2.71-8.73)	0.20 (0.02-2.53)
Perry 2006	100 (48-100)	98 (97-99)	47.67 (26.67-85.20)	0.08 (0.01-1.21)
Gangloff 2015	100 (59-100)	95 (93-96)	15.23 (1.58-146.73)	0.13 (0.02-0.83)
Pooled Accuracy				
(anthochromia Visual				
Perry 2006	50 (1-99)	97 (93-99)	15.57 (3.25-74.54)	0.52 (0.13-2.07)
Dupont 2008	93 (66-100)	95 (89-98)	19.13 (8.04-45.45)	0.08 (0.01-0.50)
Gangloff 2015	80 (28-99)	99 (98-99)	62.31 (28.47-136.37)	0.20 (0.04-1.17)
Pooled Accuracy	31 (21-41)	98 (97-99)	28.79 (9.77-84.80)	0.22 (0.06-0.80)
(anthochromia				
raditional				
Perry 2006	100 (16-100)	29 (23-35)	1.17 (0.70-1.96)	0.57 (0.05-7.28)
(anthochromia –				
Chalmers/Kiley				
Perry 2006	0 (0-84)	89 (84-93)	1.49 (0.12-19.23)	0.94 (0.56-1.56)
(anthochromia – Chalmers revised				
	100 (16-100)	29 (23-35)	1.17 (0.70-1.96)	0.57 (0.05-7.28)
Perry 2006				

### Table 6

# **Using Test-Indication Curves**

You have just seen a 48-year-old housewife who presents to the ED by private vehicle 8 hours after the onset of a severe headache which peaked within 10 minutes of onset. She is nauseated and complains of neck stiffness but her neurological exam is intact and she is able to fully flex her neck. You assign a pre-test probability of 10% for SAH, and await the final read by the neuroradiologist on duty even though the imaging study appears to be unremarkable to you and to the CT technician. The test treatment curves can be used like the Fagan nomogram to calculate the post-test probability of disease. Therefore you can estimate that the post-test probability will be around 0.8% should the study prove negative (Fig TIC panel B). You can also estimate that, if you had performed an LP rather than CT, the post-test likelihood would be around 2% if fewer than 1000 x 10<sup>6</sup>/L RBC were present in the final tube of CSF (panel C), and similarly 2% if there were no visible xanthochromia. But you can also "chain" sequential test results, with the usual caveat that the tests may not be perfectly independent. Thus, after the neuroradiologist confirms that there are no signs of SAH on the CT, the same patient would have her probability of disease lowered from 0.8% pre-LP to about 0.1% after an LP negative for xanthochromia (panel D). Moreover, decision analysis would suggest that, for this patient, the risks of LP outweigh the potential benefits, since the pre-test probability of 0.8% is already below the range when LP is beneficial whether one uses erythrocyte count or xanthochromia, as indicated by the horizontal arrows on the test indication curves.

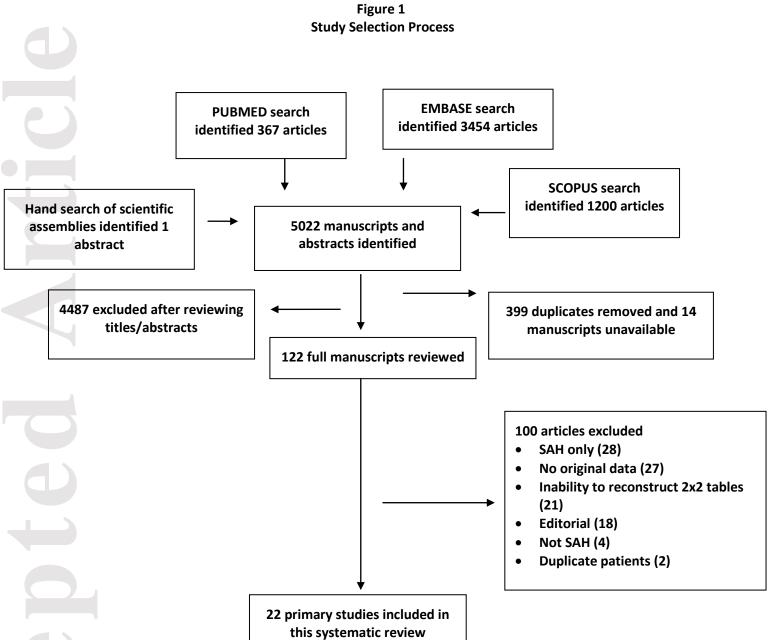
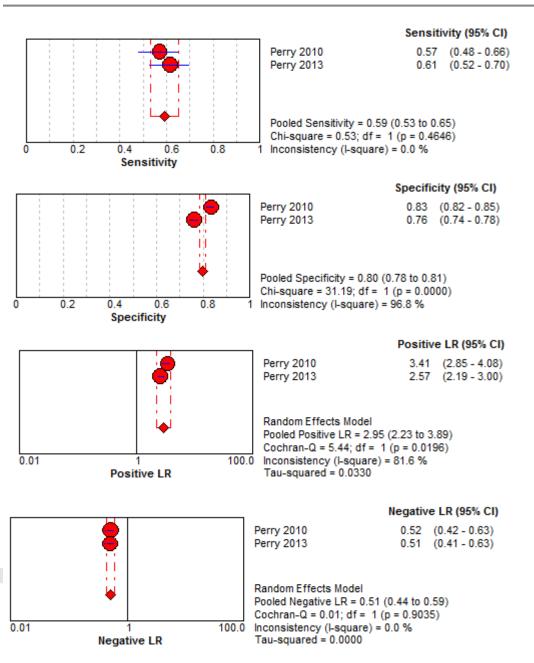
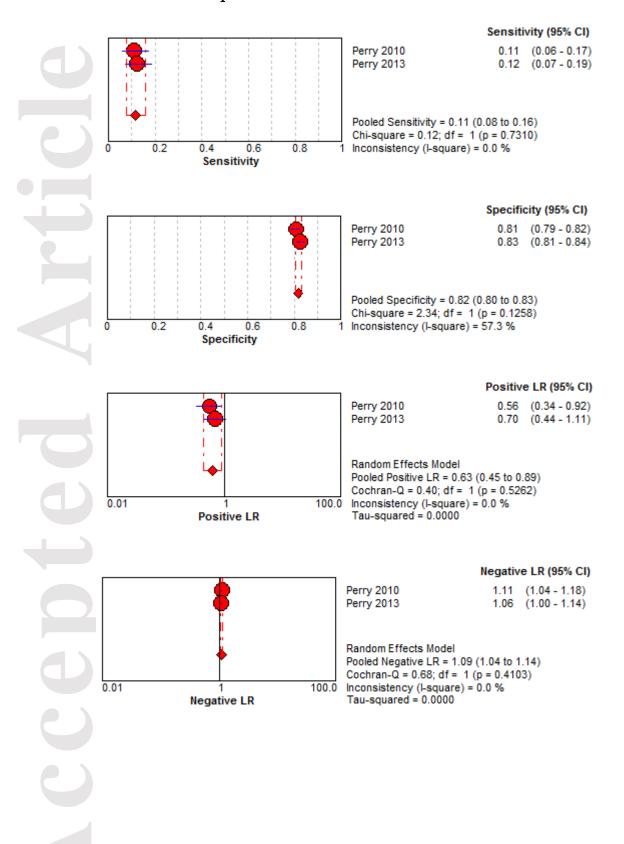


Figure 2: Forest Plots for Diagnostic Elements of History for SAH

# **Ambulance Arrival**



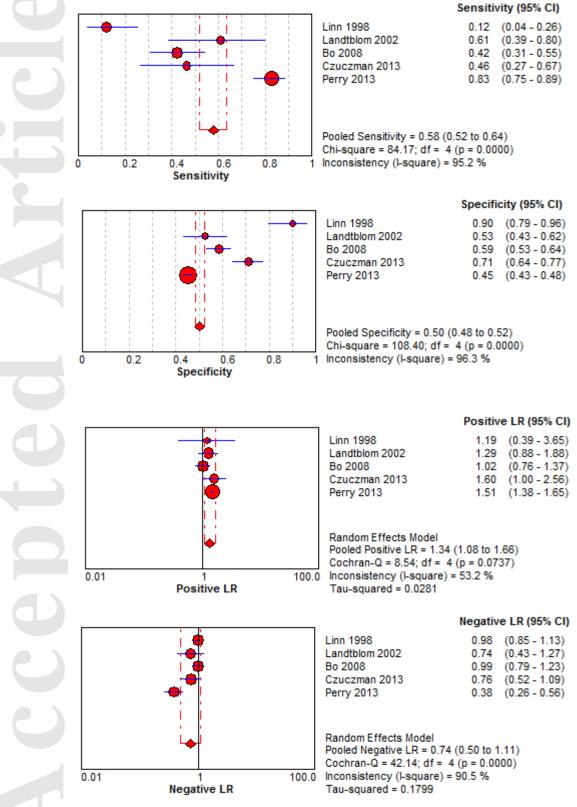
# Awoke from Sleep



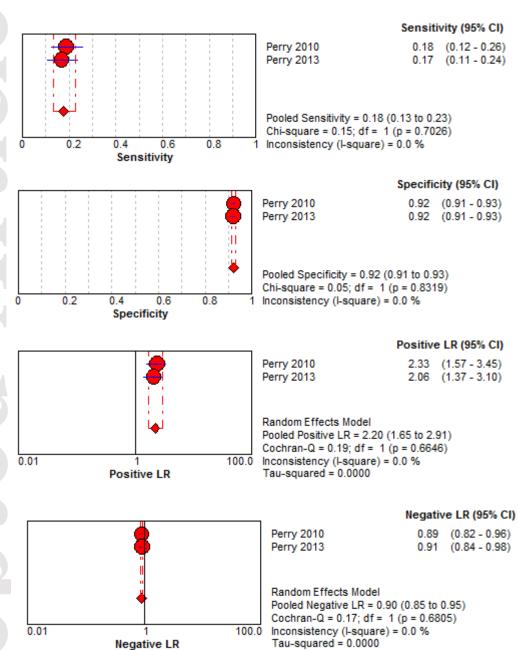
**Blurred Vision** 

# Sensitivity (95% CI) 0.04 (0.00 - 0.22) Landblom 2002 Carstairs 2006 0.40 (0.05 - 0.85) Pooled Sensitivity = 0.11 (0.02 to 0.28) Chi-square = 4.11; df = 1 (p = 0.0426) 1 Inconsistency (I-square) = 75.7 % 0.2 0.8 Sensitivity Specificity (95% CI) Landblom 2002 0.96 (0.90 - 0.99) Carstairs 2006 0.95 (0.89 - 0.98) Pooled Specificity = 0.95 (0.92 to 0.98) Chi-square = 0.08; df = 1 (p = 0.7797) 0.2 Inconsistency (I-square) = 0.0 % 0.6 0.8 Specificity Positive LR (95% CI) Landblom 2002 0.99 (0.12 - 8.09) Carstairs 2006 7.73 (2.05 - 29.14) Random Effects Model Pooled Positive LR = 3.14 (0.31 to 31.43) Cochran-Q = 3.49; df = 1 (p = 0.0617) 100.0 Inconsistency (I-square) = 71.4 % Tau-squared = 2.0010 0.01 Positive LR Negative LR (95% CI) 1.00 (0.91 - 1.10) Landblom 2002 0.63 (0.31 - 1.30) Carstairs 2006 Random Effects Model Pooled Negative LR = 0.85 (0.44 to 1.63) Cochran-Q = 3.52; df = 1 (p = 0.0608)0.01 100.0 Inconsistency (I-square) = 71.6 % Tau-squared = 0.1713 Negative LR

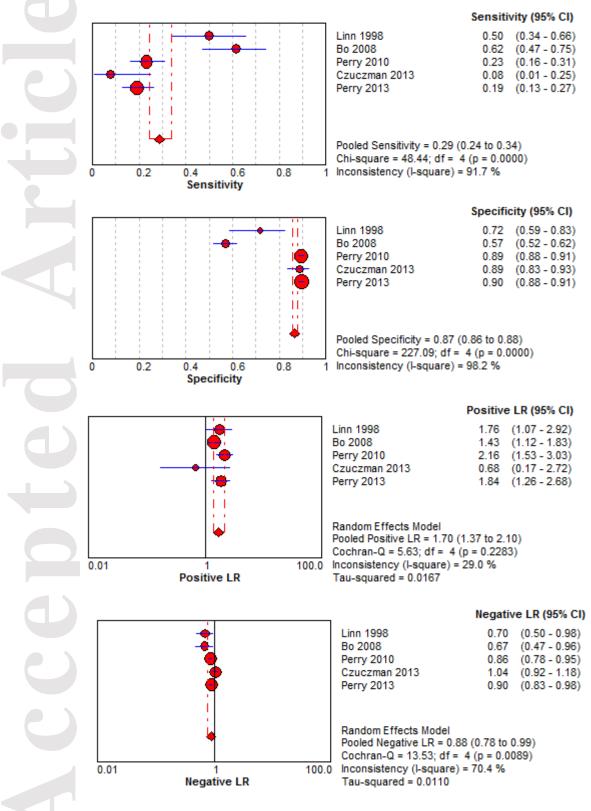
# Burst or Explode at Onset



**ED Transfer** 



# Exertion at Onset



### **Female** Sensitivity (95% CI) Linn 1998 0.57 (0.41 - 0.72) Morgenstern 1998 0.50 (0.26 - 0.74)(0.28 - 0.99)Carstairs 2006 0.80 Bo 2008 0.58 (0.45 - 0.69)Perry 2010 0.57 (0.48 - 0.66) Perry 2013 0.58 (0.49 - 0.67) Pooled Sensitivity = 0.58 (0.53 to 0.62) Chi-square = 1.60; df = 5 (p = 0.9009) 0.2 0.4 0.6 0.8 Inconsistency (I-square) = 0.0 % Sensitivity Specificity (95% CI) Linn 1998 0.68 (0.55 - 0.80) Morgenstern 1998 0.71 (0.60 - 0.81) Carstairs 2006 0.42 (0.33 - 0.52) Bo 2008 0.41 (0.36 - 0.47) Perry 2010 0.39 (0.37 - 0.42) Perry 2013 0.39 (0.37 - 0.42) Pooled Specificity = 0.41 (0.39 to 0.42) Chi-square = 51.15; df = 5 (p = 0.0000) 0.2 0.6 0.8 Inconsistency (I-square) = 90.2 % Specificity Positive LR (95% CI) Linn 1998 1.80 (1.15 - 2.84) Morgenstern 1998 1.72 (0.97 - 3.06) Carstairs 2006 (0.87 - 2.21)1.39 Bo 2008 0.99 (0.79 - 1.23) Perry 2010 0.94 (0.81 - 1.10) Perry 2013 0.96 (0.83 - 1.12) Random Effects Model Pooled Positive LR = 1.10 (0.93 to 1.31) Cochran-Q = 12.63; df = 5 (p = 0.0271) 100.0 0.01 Inconsistency (I-square) = 60.4 % Positive LR Tau-squared = 0.0232 Negative LR (95% CI) 0.63 (0.42 - 0.93) Linn 1998 Morgenstern 1998 0.71 (0.44 - 1.14) Carstairs 2006 0.47 (0.08 - 2.77) Bo 2008 1.02 (0.76 - 1.37) Perry 2010 1.09 (0.89 - 1.34) Perry 2013 1.06 (0.86 - 1.31) Random Effects Model Pooled Negative LR = 0.93 (0.78 to 1.12) Cochran-Q = 9.18; df = 5 (p = 0.1021) 0.01 100.0 Inconsistency (I-square) = 45.5 %

Tau-squared = 0.0215

Negative LR

Intercourse at Onset

### Sensitivity (95% CI) Landtblom 2003 0.09 (0.01 - 0.28) Bo 2008 0.03 (0.00 - 0.10) 0.05 (0.02 - 0.11) Perry 2010 Perry 2013 0.10 (0.05 - 0.16) Pooled Sensitivity = 0.07 (0.04 to 0.10) Chi-square = 4.52; df = 3 (p = 0.2101) Inconsistency (I-square) = 33.7 % 0.2 0.4 0.6 0.8 Sensitivity Specificity (95% CI) Landtblom 2003 0.93 (0.88 - 0.97) Bo 2008 0.94 (0.92 - 0.97) Perry 2010 0.94 (0.93 - 0.95) Perry 2013 0.94 (0.93 - 0.95) Pooled Specificity = 0.94 (0.93 to 0.95) Chi-square = 0.33; df = 3 (p = 0.9535) 1 Inconsistency (I-square) = 0.0 % 0.2 0.6 0.8 Specificity Positive LR (95% CI) Landtblom 2003 1.32 (0.31 - 5.74) Bo 2008 0.51 (0.12 - 2.13) Perry 2010 0.90 (0.43 - 1.89) Perry 2013 1.59 (0.92 - 2.73) Random Effects Model Pooled Positive LR = 1.20 (0.79 to 1.82) Cochran-Q = 3.10; df = 3 (p = 0.3772) 100.0 Inconsistency (I-square) = 3.1 % 0.01 Positive LR Tau-squared = 0.0067 Negative LR (95% CI) Landtblom 2003 0.98 (0.85 - 1.12) Bo 2008 1.03 (0.98 - 1.08) Perry 2010 1.01 (0.96 - 1.05) Perry 2013 0.96 (0.91 - 1.02) Random Effects Model Pooled Negative LR = 1.00 (0.97 to 1.03) Cochran-Q = 4.02; df = 3 (p = 0.2594) 100.0 Inconsistency (I-square) = 25.4 % 0.01 Tau-squared = 0.0003 Negative LR

### Loss of Consciousness Sensitivity (95% CI) Linn 1998 0.26 (0.14 - 0.42) Landtblom 2002 0.17 (0.05 - 0.39) 0.17 (0.11 - 0.24) Perry 2010 0.11 (0.06 - 0.17) Perry 2013 Pooled Sensitivity = 0.16 (0.12 to 0.20) Chi-square = 6.07; df = 3 (p = 0.1082) 1 Inconsistency (I-square) = 50.6 % 0 0.2 0.4 0.6 0.8 Specificity (95% CI) Linn 1998 0.42 (0.15 - 0.72) Landtblom 2002 0.96 (0.90 - 0.99) Perry 2010 0.96 (0.94 - 0.96) 0.95 (0.94 - 0.96) Perry 2013 Pooled Specificity = 0.95 (0.94 to 0.96) Chi-square = 27.64; df = 3 (p = 0.0000) Inconsistency (I-square) = 89.1 % 0.2 0.6 Specificity Positive LR (95% CI) Linn 1998 0.45 (0.22 - 0.90) 3.97 (1.15 - 13.65) Landtblom 2002 Perry 2010 3.77 (2.44 - 5.81) 2.00 (1.18 - 3.39) Perry 2013 Random Effects Model Pooled Positive LR = 1.87 (0.72 to 4.86) Cochran-Q = 26.78; df = 3 (p = 0.0000) 0.01 100.0 Inconsistency (I-square) = 88.8 % Positive LR Tau-squared = 0.8086 Negative LR (95% CI) Linn 1998 1.77 (0.89 - 3.54) Landtblom 2002 0.86 (0.71 - 1.05) Perry 2010 0.87 (0.80 - 0.94) Perry 2013 0.94 (0.89 - 1.00) Random Effects Model Pooled Negative LR = 0.91 (0.83 to 1.00) Cochran-Q = 6.49; df = 3 (p = 0.0899)Inconsistency (I-square) = 53.8 % 100.0 0.01

Tau-squared = 0.0037

Negative LR

0.01

### Male Sensitivity (95% CI) Morgantern 1998 0.50 (0.26 - 0.74) 0.20 (0.01 - 0.72) Carstairs 2006 Bo 2008 0.42 (0.31 - 0.55) Pooled Sensitivity = 0.43 (0.32 to 0.53) Chi-square = 1.55; df = 2 (p = 0.4617)0.2 0.4 0.6 0.8 Inconsistency (I-square) = 0.0 % Sensitivity Specificity (95% CI) 0.29 (0.19 - 0.40) Morgantern 1998 0.58 (0.48 - 0.67) Carstairs 2006 Bo 2008 0.59 (0.53 - 0.64) Pooled Specificity = 0.54 (0.50 to 0.58) Chi-square = 23.72; df = 2 (p = 0.0000) 0.2 1 Inconsistency (I-square) = 91.6 % 0.6 0.8 Specificity Positive LR (95% CI) Morgantern 1998 0.71 (0.44 - 1.14) 0.47 (0.08 - 2.77) Carstairs 2006 Bo 2008 1.02 (0.76 - 1.37) Random Effects Model Pooled Positive LR = 0.90 (0.68 to 1.18) Cochran-Q = 2.16; df = 2 (p = 0.3400) 100.0 Inconsistency (I-square) = 7.3 % 0.01 Positive LR Tau-squared = 0.0061 Negative LR (95% CI) 1.72 (0.97 - 3.06) Morgantern 1998 Carstairs 2006 1.39 (0.87 - 2.21) Bo 2008 0.99 (0.79 - 1.23) Random Effects Model Pooled Negative LR = 1.23 (0.88 to 1.72) Cochran-Q = 4.24; df = 2 (p = 0.1201)

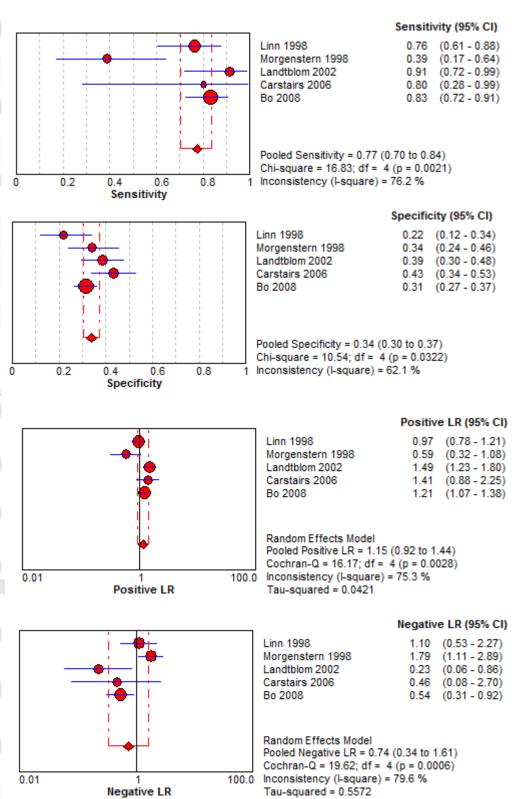
100.0

Inconsistency (I-square) = 52.8 %

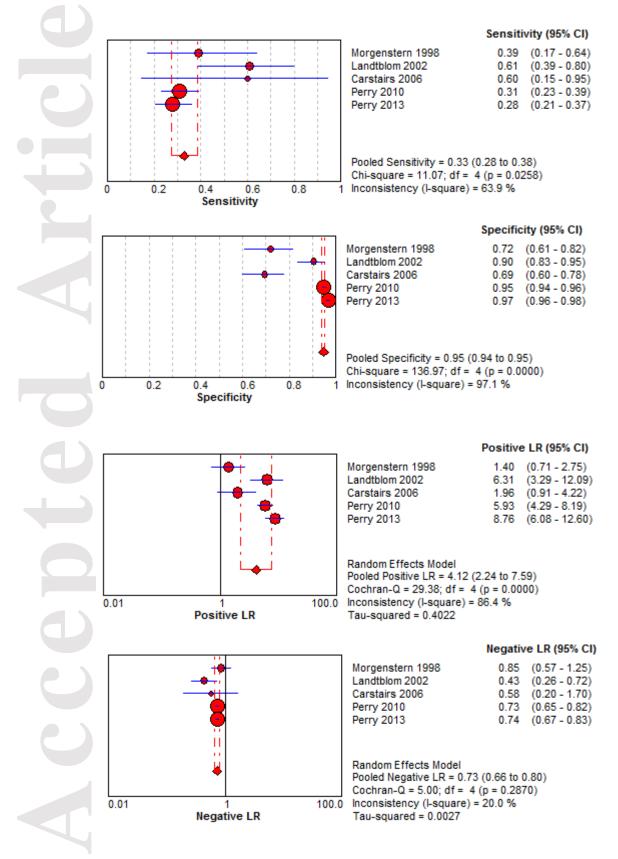
Tau-squared = 0.0478

Negative LR

# Nausea 0.2 0.2 0.01 0.01

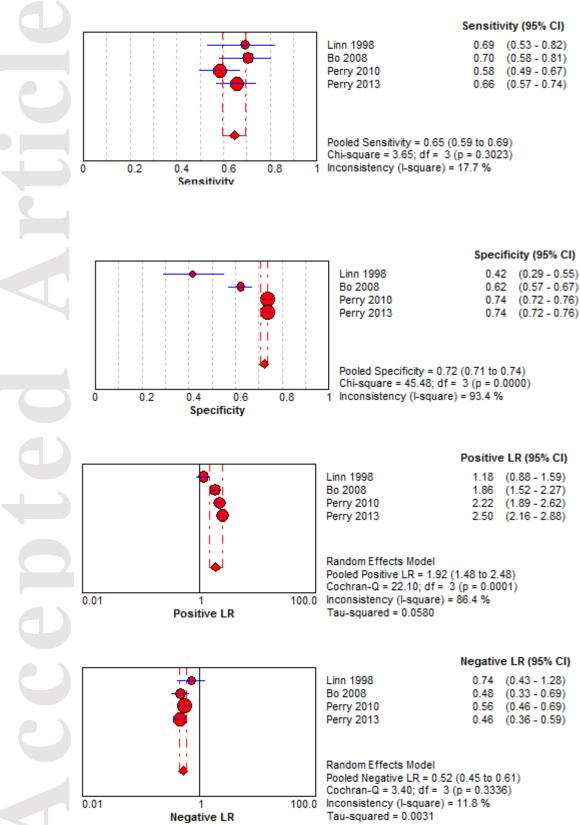


# Neck stiffness, subjective



### Photophobia Sensitivity (95% CI) 0.28 (0.10 - 0.53) Morgenstern 1998 Landtblom 2002 0.09 (0.01 - 0.28) 0.80 (0.28 - 0.99) Carstairs 2006 Bo 2008 0.46 (0.35 - 0.59) Pooled Sensitivity = 0.38 (0.29 to 0.47) Chi-square = 16.99; df = 3 (p = 0.0007) 0.2 0.4 0.6 8.0 1 Inconsistency (I-square) = 82.3 % Sensitivity Specificity (95% CI) Morgenstern 1998 0.54 (0.43 - 0.66) Landtblom 2002 0.96 (0.90 - 0.99) 0.51 (0.42 - 0.61) Carstairs 2006 Bo 2008 0.49 (0.43 - 0.54) Pooled Specificity = 0.58 (0.54 to 0.62) Chi-square = 101.66; df = 3 (p = 0.0000) 1 Inconsistency (I-square) = 97.0 % 0.6 Specificity Positive LR (95% CI) Morgenstern 1998 0.61 (0.28 - 1.33) 1.98 (0.41 - 9.60) Landtblom 2002 Carstairs 2006 1.64 (1.02 - 2.65) Bo 2008 0.90 (0.69 - 1.18) Random Effects Model Pooled Positive LR = 1.07 (0.67 to 1.71) Cochran-Q = 7.64; df = 3 (p = 0.0540) 0.01 100.0 Inconsistency (I-square) = 60.7 % Positive LR Tau-squared = 0.1246 Negative LR (95% CI) 1.33 (0.93 - 1.88) Morgenstern 1998 Landtblom 2002 0.95 (0.84 - 1.09) 0.39 (0.07 - 2.27) Carstairs 2006 Bo 2008 1.10 (0.86 - 1.40) Random Effects Model Pooled Negative LR = 1.05 (0.87 to 1.27) Cochran-Q = 5.19; df = 3 (p = 0.1584) 0.01 100.0 Inconsistency (I-square) = 42.2 % Tau-squared = 0.0146 Negative LR

# Vomiting



# Worst Headache of Life

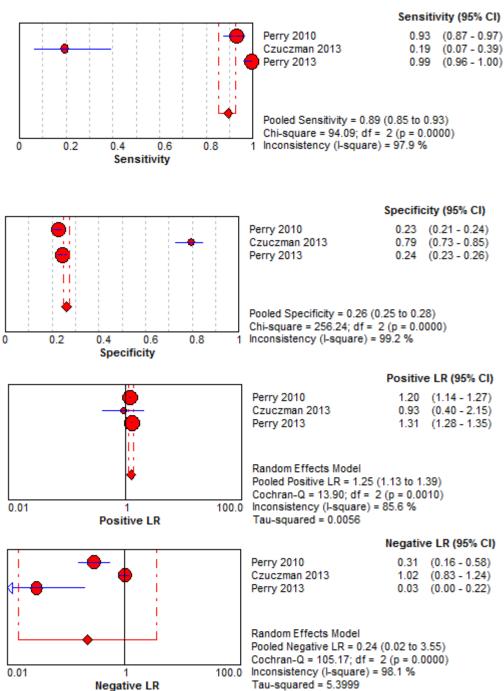
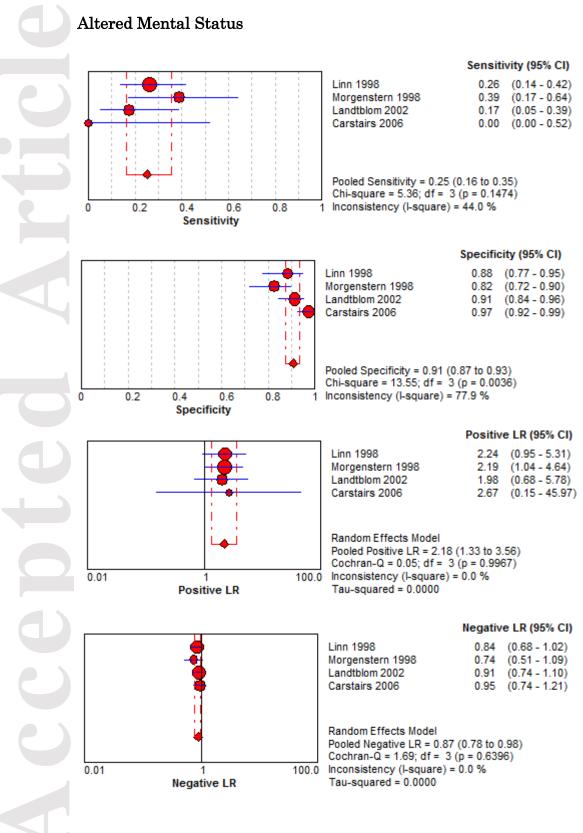


Figure 3: Forest Plots for Diagnostic Elements of Physical Exam for SAH



### Focal Neuro Deficit Sensitivity (95% CI) Linn 1998 0.33 (0.20 - 0.50) Morgenstern 1998 0.50 (0.26 - 0.74) Landtblom 2002 0.13 (0.03 - 0.34) Carstairs 2006 0.20 (0.01 - 0.72) Pooled Sensitivity = 0.31 (0.21 to 0.41) Chi-square = 7.27; df = 3 (p = 0.0637) Inconsistency (I-square) = 58.8 % 0.2 0.4 8.0 Sensitivity Specificity (95% CI) Linn 1998 0.83 (0.71 - 0.92) Morgenstern 1998 0.90 (0.81 - 0.96)Landtblom 2002 0.97 (0.93 - 0.99)Carstairs 2006 0.95 (0.90 - 0.99) Pooled Specificity = 0.93 (0.90 to 0.95) Chi-square = 12.94; df = 3 (p = 0.0048) Inconsistency (I-square) = 76.8 % 0.2 0.6 8.0 Specificity Positive LR (95% CI) Linn 1998 2.00 (0.98 - 4.06) 4.94 (2.21 - 11.02) Morgenstern 1998 4.96 (1.07 - 23.04) Landtblom 2002 Carstairs 2006 4.44 (0.63 - 31.24) Random Effects Model Pooled Positive LR = 3.26 (1.93 to 5.52) Cochran-Q = 3.29; df = 3 (p = 0.3486) 0.01 100.0 Inconsistency (I-square) = 8.9 % Positive LR Tau-squared = 0.0287 Negative LR (95% CI) Linn 1998 0.80 (0.63 - 1.02) Morgenstern 1998 0.56 (0.35 - 0.89) Landtblom 2002 0.89 (0.76 - 1.05) Carstairs 2006 0.84 (0.54 - 1.30) Random Effects Model Pooled Negative LR = 0.81 (0.67 to 0.97) Cochran-Q = 4.99; df = 3 (p = 0.1726)100.0 Inconsistency (I-square) = 39.9 % 0.01

Tau-squared = 0.0138

Negative LR

# Neck stiffness, objective

# Sensitivity (95% CI) Carstairs 2006 0.00 (0.00 - 0.60) Perry 2010 0.31 (0.23 - 0.39) Perry 2013 0.28 (0.21 - 0.37) Pooled Sensitivity = 0.29 (0.24 to 0.35) Chi-square = 3.00; df = 2 (p = 0.2236) 0.2 Inconsistency (I-square) = 33.2 % Sensitivity Specificity (95% CI) 0.85 (0.77 - 0.91) Carstairs 2006 Perry 2010 0.95 (0.94 - 0.96) Perry 2013 0.97 (0.96 - 0.98) Pooled Specificity = 0.96 (0.95 to 0.96) Chi-square = 29.70; df = 2 (p = 0.0000) Inconsistency (I-square) = 93.3 % 0.2 0.6 0.8 Specificity Positive LR (95% CI) Carstairs 2006 0.64 (0.04 - 9.19) Perry 2010 5.93 (4.29 - 8.19) Perry 2013 8.76 (6.08 - 12.60) Random Effects Model Pooled Positive LR = 6.59 (3.95 to 11.00) Cochran-Q = 5.69; df = 2 (p = 0.0581) Inconsistency (I-square) = 64.9 % 100.0 0.01 Positive LR Tau-squared = 0.1111

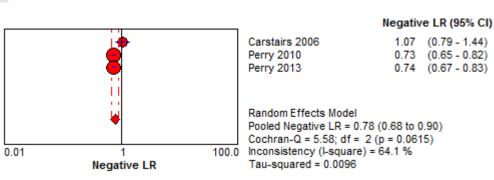


Figure 4: Forest Plots for Diagnostic Accuracy of CT for SAH

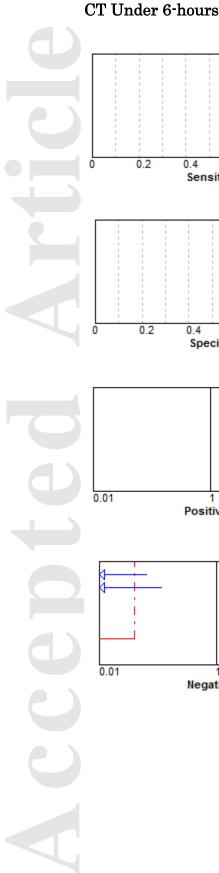
### **CT** Anytime Sensitivity (95% CI) VanderWee 1994 0.98 (0.94 - 1.00) Boesiger 2005 1.00 (0.54 - 1.00) O'Neill 2005 0.76 (0.55 - 0.91) Perry 2011 0.93 (0.89 - 0.96)Backes 2012 0.95 (0.89 - 0.98) Pooled Sensitivity = 0.94 (0.91 to 0.96) Chi-square = 15.50; df = 4 (p = 0.0038)8.0 Inconsistency (I-square) = 74.2 % 0.2 0.6 Sensitivity Specificity (95% CI) VanderWee 1994 1.00 (0.94 - 1.00) Boesiger 2005 0.99 (0.97 - 1.00) O'Neill 2005 1.00 (0.96 - 1.00) Perry 2011 1.00 (1.00 - 1.00) Backes 2012 1.00 (0.97 - 1.00) Pooled Specificity = 1.00 (1.00 to 1.00) Chi-square = 5.96; df = 4 (p = 0.2023) 0.2 0.6 0.8 Inconsistency (I-square) = 32.9 % Specificity Positive LR (95% CI) VanderWee 1994 111.63 (7.07 - 1,763.23) Boesiger 2005 106.48 (21.36 - 530.82) O'Neill 2005 138.00 (8.62 - 2,209.28) 5,365.85 (335.63 - 85,786.14) Perry 2011 Backes 2012 276.85 (17.39 - 4,406.77) Random Effects Model Pooled Positive LR = 267.26 (57.67 to 1238.49) Cochran-Q = 7.72; df = 4 (p = 0.1022) 100.0 Inconsistency (I-square) = 48.2 % 0.01 Positive LR Tau-squared = 1.4480 Negative LR (95% CI) 0.02 (0.01 - 0.07) VanderWee 1994 Boesiger 2005 0.07 (0.00 - 1.04) O'Neill 2005 0.25 (0.13 - 0.49) Perry 2011 0.07 (0.05 - 0.11) Backes 2012 0.05 (0.02 - 0.12) Random Effects Model Pooled Negative LR = 0.07 (0.03 to 0.17) Cochran-Q = 17.76; df = 4 (p = 0.0014) 0.01 100.0 Inconsistency (I-square) = 77.5 % Negative LR Tau-squared = 0.5862

CT Over 6-hours

## Sensitivity (95% CI) Perry 2011 0.86 (0.78 - 0.91) Backes 2012 1.00 (0.89 - 1.00) Pooled Sensitivity = 0.89 (0.83 to 0.93) Chi-square = 8.90; df = 1 (p = 0.0029) Inconsistency (I-square) = 88.8 % Sensitivity Specificity (95% CI) Perry 2011 1.00 (1.00 - 1.00) Backes 2012 0.95 (0.88 - 0.99) Pooled Specificity = 1.00 (1.00 to 1.00) Chi-square = 26.49; df = 1 (p = 0.0000) 0.2 0.6 0.8 Inconsistency (I-square) = 96.2 % Specificity Positive LR (95% CI) 3,520.87 (220.08 - 56,327.07) Perry 2011 Backes 2012 17.74 (7.22 - 43.57) Random Effects Model Pooled Positive LR = 223.37 (0.39 to 129409.88) Cochran-Q = 19.09; df = 1 (p = 0.0000)100.0 0.01 Inconsistency (I-square) = 94.8 % Positive LR Tau-squared = 20.0045 Negative LR (95% CI) Perry 2011 0.15 (0.09 - 0.22) 0.02 (0.00 - 0.24) Backes 2012 Random Effects Model Pooled Negative LR = 0.07 (0.01 to 0.61) Cochran-Q = 2.72; df = 1 (p = 0.0991)0.01 Inconsistency (I-square) = 63.2 %

Tau-squared = 1.7372

Negative LR



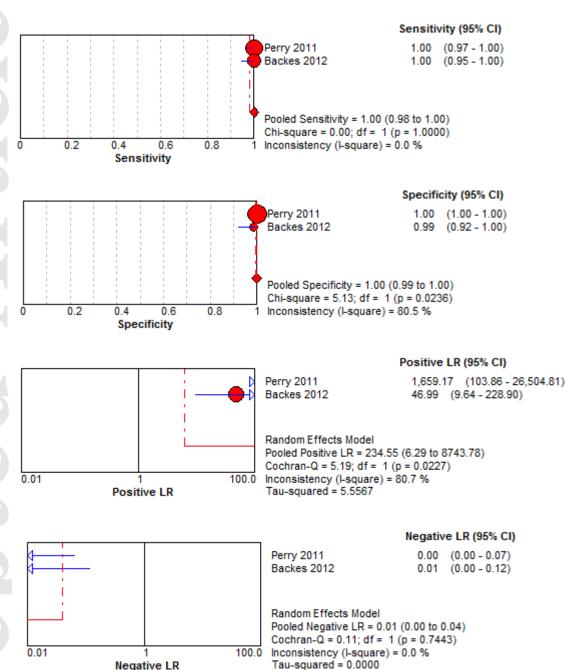
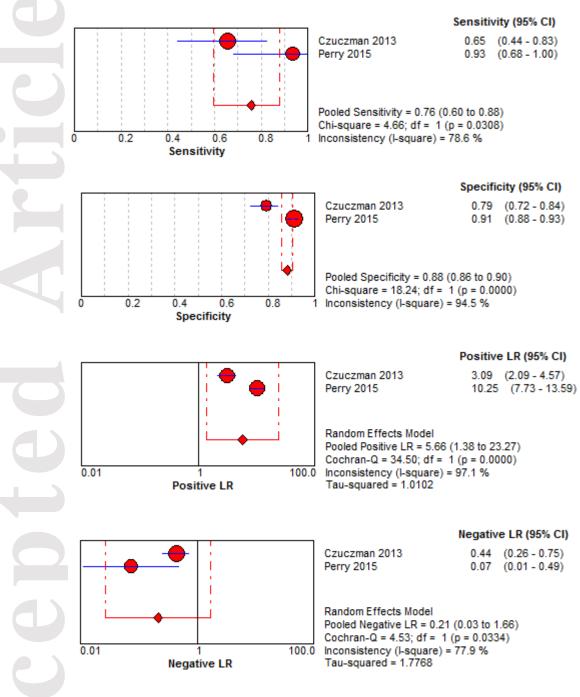
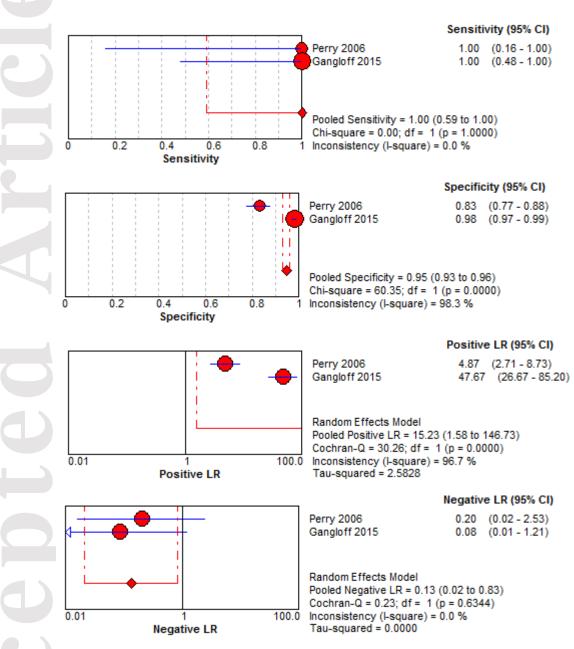


Figure 5: Forest Plots for Diagnostic Accuracy of CSF Analysis for SAH CSF RBC >1000 x 106/L



## LP--UK Spectrophometric Xanthochromia



## LP--Visible Xanthochromia

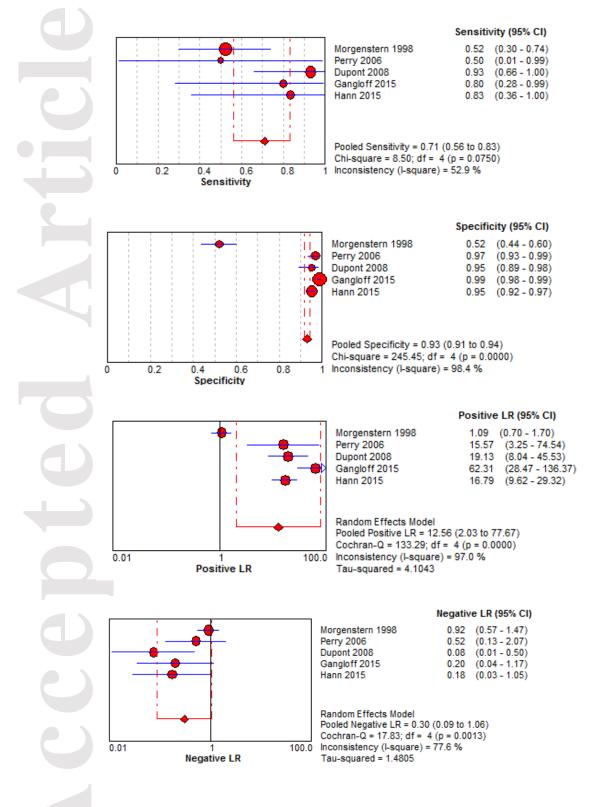
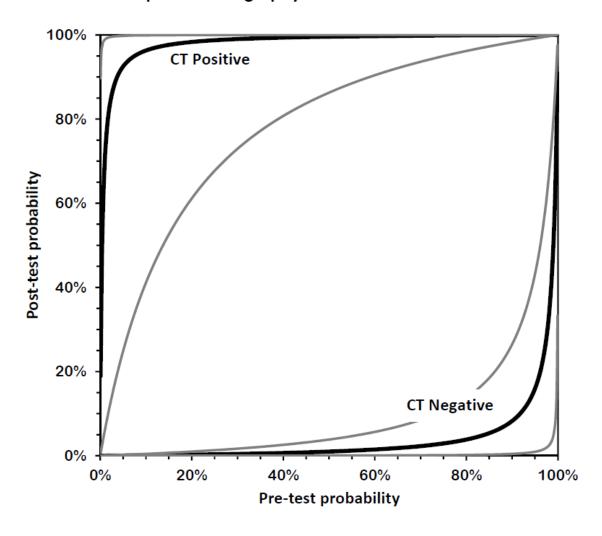


Figure 6: Bernstein Test-Indication Curves for CT and LP

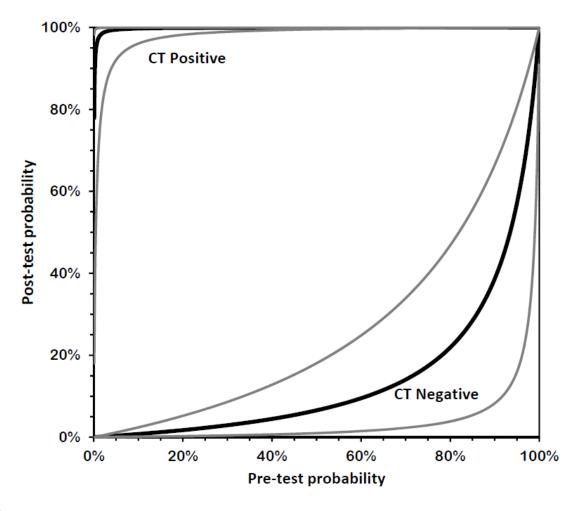
A: CT Within 6 Hours of Headache Onset

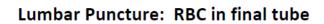
# Unenhanced Computed Tomography Within 6 hours of Headache Onset

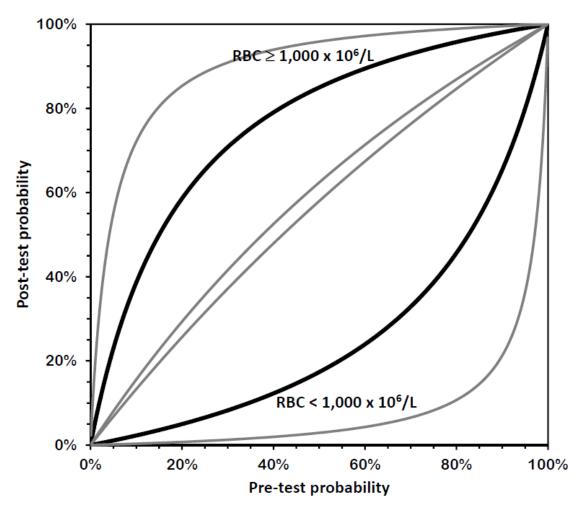


B: CT Beyond 6 Hours of Headache Onset

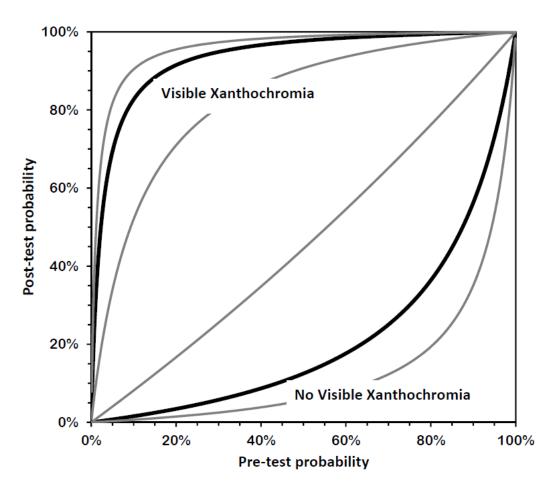
## Unenhanced Computed Tomography Beyond 6 hours of Headache Onset







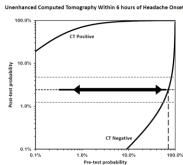
## Lumbar Puncture: Xanthochromia



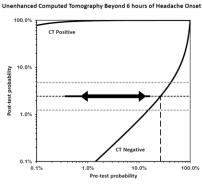
Test Indication Curves with 95% CI natural scale in panels A through D. **Diagnostic** accuracy of computed tomography or lumbar puncture for subarachnoid hemorrhage. Raw test-indication curves as proposed by Bernstein are shown, providing a graphical representation of the Bayesian post-test probability (y-axis) based on either a positive (upper curved black line, with 95% confidence intervals in grey) or negative (lower curved lines) test result as a function of the pre-test probability (x-axis). The graphs use the same principle as the Fagan nomogram, but provide more intuitive representations of the diagnostic accuracy of a test. The four tests considered are cranial CT obtained within 6 hours (panel A) or later (panel B) from headache onset, or LP with more than 1,000x10<sup>6</sup>/L erythrocytes (panel C) or visible xanthochromia (panel D). The distance of the curves from the main diagonal of zero diagnostic information provide a visual representation of the information gained from the test result.

Figure 7 Testing Thresholds

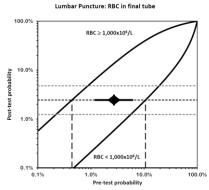
## A: CT Within 6 Hours of Headache Onset



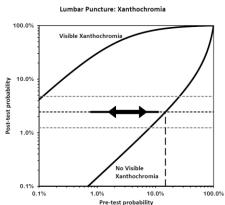
## B: CT Beyond 6 Hours of Headache Onset



## C: LP Using RBC Counts in the Final Tube



### D: LP Using Visible Xanthochromia



Test-indication curves illustrating when benefits of testing outweigh the risks for **subarachnoid hemorrhage.** The complete test-indication curves as proposed by Bernstein are shown which incorporate the Pauker-Kassirer threshold approach to deciding whether to perform the test in question. This technique compares the risks versus benefits for treatment (in this case proceeding to CT or formal angiography), and calculates the corresponding treatment threshold (shown as horizontal dashed lines, with the point estimate in black and the upper and lower bands of the sensitivity analysis in grey). The vertical line dashed line (or rightmost line when two are visible) therefore represents the pre-test probability range below which the diagnostic test might be appropriate, assuming the test in question had neither risks nor costs; when the pre-test probability is higher, empirical treatment is recommended since the post-test probability exceeds this threshold even if the test is negative. Because tests have a non-zero risk, the actual pre-test range for which performing the test is rational is narrower, and is shown by the thick line between the arrowheads. The thinner solid line extending beyond the arrows illustrates the effect of reducing the test risk by half. At pre-test probabilities to the left of this range, no further testing is indicated. For example, this approach suggests proceeding directly to angiography when the pre-CT probability exceeds 10% (panel B), unless the unenhanced CT can be obtained within 6 hours (panel A) in which case a pre-test probability of nearly 70% seems appropriate before proceeding directly to angiography. On the other hand, the use of CT is warranted even in very low risk patients, down to perhaps 0.7% pre-test probability as determined primarily by the risks of the CT itself. The LP curves, however, illustrate that the pre-LP probabilities that justify performing LP are very narrow, ranging from 2% to 4% in panel C and 2% to 7% in panel D. Moreover, these pre-LP probabilities in turn can only arise after an implausibly high pre-CT probability (>>20%).