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Determination of a Testing Threshold for Lumbar Puncture in the Diagnosis of Subarachnoid Hemorrhage after a Negative Head CT: A Decision Analysis

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Author Contributions: 2

RAT and DHN conceived, designed, and supervised the study. HSG and EGM provided expert review of the structure of the decision analysis. HPM, JSF, and DHN undertook literature review and acquisition of the data. RAT analyzed the data. RAT drafted the manuscript, and all authors contributed substantially to its revision. RAT takes responsibility for the paper as a whole.

Abstract:

Objective

To determine the testing threshold for lumbar puncture (LP) in the evaluation of aneurysmal subarachnoid hemorrhage (SAH) after a negative head CT. As a secondary aim we sought to identify clinical variables that have the greatest impact on this threshold.

Methods

A decision analytic model was developed to estimate the testing threshold for patients with normal neurologic findings, being evaluated for SAH, after a negative CT of the head. The testing threshold was calculated as the pretest probability of disease where the two strategies (LP or no LP) are balanced in terms of quality adjusted life years (QALYs). Two-way and probabilistic sensitivity analyses (PSA) were performed.

Results

For the base case scenario the testing threshold for performing an LP after negative head CT was 4.3%. Results for the two-way sensitivity analyses demonstrated that the test threshold ranged from 1.9%-15.6%, dominated by the uncertainty in the probability of death from initial missed SAH. In the PSA the mean testing threshold was 4.3% (95%CI, 1.4-9.3). Other significant variables in the model included: probability of aneurysmal versus non-aneurysmal SAH after negative head CT, probability of long-term morbidity from initial missed SAH, and probability of renal failure from contrast induced nephropathy.

Conclusions

Our decision analysis results suggest a testing threshold for LP after negative CT to be approximately 4.3%, with a range of 1.4% to 9.3% on robust PSA. In light of these data, and considering the low probability of aneurysmal SAH after a negative CT, classical teaching and current guidelines addressing testing for subarachnoid hemorrhage should be revisited.

Introduction

Background

Aneurysmal sub-arachnoid hemorrhage (SAH) is a common concern in the evaluation of neurologically normal patients with headache, but an uncommon occurrence. Headaches account for approximately 2% of annual emergency department (ED) visits, though SAH accounts for less than 1% of these.¹ Misdiagnosis and morbidity rates associated with SAH are high,² and current clinical practice guidelines^{3,4} recommend lumbar puncture after a negative non-contrast computed tomography (CT) despite a very low probability of disease.⁵

Importance

Lumbar puncture, however, is not without risks including meningitis, neurologic injury, and patient harm from further pursuit of false positive results.⁶⁻⁸ Attempts to balance these complex processes has made the decision of whether to perform lumbar puncture following CT a decision point with both high clinician variability^{2,9} and with the potential for important impact on patient outcomes.¹⁰

Decision analysis is a mathematical modeling technique well suited to analyzing complex medical problems with multiple components and determining optimal decision strategies under varying conditions.¹¹ Among headache patients considered for lumbar puncture, decision analysis allows for the determination of the comparative impact, in quality adjusted life years (QALYs), of performing versus not performing lumbar puncture at different pretest probabilities of disease. The pretest probability of disease where the two strategies are balanced in terms of QALYs is known as

the testing threshold, and represents the acceptable miss rate (i.e. if testing was performed at pretest probabilities lower than the threshold the risks of harm to the patient from further testing would outweigh the risk of benefit).¹²

Goals of This Investigation

The primary aim of our study was to determine the testing threshold for lumbar puncture in the evaluation of SAH after a negative non-contrast head CT. As a secondary aim we planned to identify clinical variables that have the greatest impact on this threshold.

Materials and Methods Study Design

This study was a decision analysis developed according to published guidelines¹³ to estimate the testing threshold for performing lumbar puncture after a negative non-contrast CT in the evaluation of a patient for SAH. As a decision analysis, the study was only dependent upon data from literature review or expert opinion and was exempt from review by our institutional review board. **Setting**

The hypothetical base case for our decision analytic model is a 45-year-old patient presenting to the emergency department with a headache and normal neurologic findings, being evaluated for SAH, after a negative non-contrast computed tomography of the head. Forty-five years represents roughly the mean age of neurologically normal patients enrolled in prospective studies designed to capture patients with SAH and those in whom SAH is an important diagnostic consideration.^{9,10,14} In our clinical scenario after negative imaging the provider is confronted with two potential diagnostic strategies: perform lumbar puncture with further testing guided by the results, or omit lumbar puncture and presumably discharge the patient. Subsequent diagnostic and management strategies were chosen to reflect standard practice and accepted guidelines for the evaluation and management of SAH.^{3,15,16}

Model Structure

To model the clinical scenario and diagnostic pathways described above, we constructed a decision tree (Figure 1) using decision analysis software (TreeAge Pro 2013, TreeAge Software, Williamstown, MA). The primary node of the decision tree represents the decision to perform or not perform lumbar puncture. If the provider chooses to perform lumbar puncture and the findings are positive for xanthochromia or blood then computed tomographic angiography is performed. If the lumbar puncture is negative, the patient is not further evaluated. For those patients with a positive lumbar puncture, if angiography is positive for a cerebrovascular aneurysm, then the patient potentially undergoes surgical or endovascular aneurysmal repair. Those patients with a positive lumbar and with a negative angiography are not further evaluated for SAH and are treated with standard care. Branch probabilities for the nodes of these diagnostic tests represent the sensitivity and specificity of the test transformed through Bayesian revision into decision probabilities (i.e. the false positive, false negative, true positive, and true negative probabilities). Additional branch points within the model represent the probability of certain events occurring (chance nodes) and the transition between several disease states (e.g. cancer) with continuing risk over time (Markov nodes).

Terminal nodes within the model represent final outcomes and were assigned values or "payoffs" based on QALYs.¹⁷ For each year within the model, a particular outcome is associated with a utility value that estimates the quality of life for that individual in a particular disease state with death equal to zero and perfect health equal to one. To account for the comparative value of future life-years we assumed a standard discount rate of 3%.¹⁸ For health states in which more than one disease state was possible (e.g. a patient having cancer and long-term morbidity from SAH) the utility values were multiplied together to obtain the composite utility value.¹⁹

Several assumptions were made in the construction of the model to decrease complexity. First, in the first year only the outcome of mortality was considered. This restriction enabled the exclusion of multiple potential branch points for short-term (<1 year) morbidity that were unlikely to have overall model effects including: effects from contrast induced nephropathy (CIN) without the need for renal replacement therapy (RRT); and short-term morbidity from RRT, lumbar puncture, anaphylaxis, and SAH. Second, we assumed that death directly attributable to meningitis from

lumbar puncture, anaphylaxis, or dialysis-dependent renal failure only occurred within the first year. Third, gender differences in outcomes were not explicitly built into the model, however gender was factored into the upper range of life expectancy. Last, we assumed the patient did not have comorbidities that would differentially affect outcomes.

Model Inputs and Data Collection

Data for the model inputs (Tables 1 & 2) were obtained from a methodical literature search and review, with ancestral search of available evidence for each topic. Using the Integrated Search Interface Web of Knowledge, Google Scholar, and PubMed in 2015, we searched for articles by combining terms *subarachnoid hemorrhage*, *lumbar puncture*, *CT or computed tomography, contrastinduced nephropathy or acute renal failure*, *radiation*, and *cerebral aneurysm* in logic based queries. Two investigators (HM and JF) blinded to study hypothesis reviewed articles or abstracts to determine relevance, extract data using a standardized data form, and grade methodologic quality according to standardized criteria, with disputes resolved by a third reviewer (DN). Credible intervals were constructed using the range of values suggested in the literature, with embedded confidence intervals where appropriate (typically from systematic reviews with high quality, low heterogeneity metaanalysis).

In modeling the clinical evaluation and treatment of SAH, there are a number of pathophysiologic factors and clinical complications that were considered as inputs for the model. Not all SAHs are the result of aneurysmal bleeding, particularly after a negative head CT. This is supported by data from the largest prospective cohort study examining emergency department headache evaluation for SAH, in which approximately 20% of patients diagnosed with SAH after negative computed tomography evaluation were found to have evidence of an underlying aneurysm.^{9,20}¹There is an estimated 0.4-7% prevalence of asymptomatic aneurysms in the general population.²¹ For false positive lumbar punctures (i.e. false positive SAH), aneurysms found on CTA will be presumed to be causative and will typically lead to therapeutic procedures (e.g. surgical clipping or endovascular coiling) with the potential for complications including death.^{22,23} There are also multiple complications that may arise from lumbar puncture. In our model post lumbar puncture headache and the discomfort of the procedure itself were not considered, as we felt these represented short-term morbidity that would be difficult to reliably model or convert into QALYs.²⁴ However, we did consider the very small risks of meningitis and paraparesis from lumbar puncture as these contribute to mortality and long-term morbidity.⁶⁻⁸ Given the limited amount of information on LP adverse outcomes, within our sensitivity analyses we set the lower bounds of mortality and morbidity rates to 0%. In addition, the model includes complications of contrast administered for CTA (e.g., death from anaphylaxis and renal failure, and long-term dialysis dependence).²⁵⁻²⁹ Furthermore, the transition of the patient through various states of cancer(i.e. no cancer, cancer, and death) and from the state of having long-term morbidity from SAH to death form SAH were incorporated into Markov nodes.

Data Analysis

To determine a testing threshold for lumbar puncture for our base case a one-way sensitivity analysis was performed to examine the impact that pretest probability of disease has on the model while other variable inputs were held constant. The testing threshold is the pretest probability at which both decisions are equally effective (i.e. produce the same number of QALYs).

Two-way sensitivity analyses were performed to evaluate the influence of model variables on the testing threshold and to account for variable uncertainty in the model. When available, a range of values for each variable was obtained from 95% confidence intervals and credible intervals constructed from literature searches as note above. When these were unavailable a range was derived through assumption and group consensus. Results of the two-way sensitivity analyses are expressed in a tornado diagram.

A limitation of two-way sensitivity analysis is that it is only two-way (i.e. all other variables except two are held constant) and thus unable to examine uncertainty within the model that results from the interaction of more than two variables. To better determine the uncertainty and range of values for the testing threshold, we further analyzed the model through probabilistic sensitivity analysis (PSA) and Monte Carlo Simulation.³⁰ We assumed beta probability distributions for the model variables with distribution parameters determined by data available from literature review or, when not available, assumption and group consensus.³¹ Monte Carlo simulation was performed with

500,000 iterations in which each iteration selected random values from the probability distributions of each variable. A testing threshold with 95% confidence intervals (CI) was determined by analyzing the values of the pretest probability of disease for iterations in which the outcomes for the two decision strategies were equal.

Results

For the base case scenario (45 year-old presenting to the ED with a headache, normal neurologic status, and negative head computed tomography) based on one-way sensitivity analysis the testing threshold for lumbar puncture was 4.3%. Adjustments for gender based life expectancy had no effect.

Results of the two-way sensitivity analyses for each variable in the model are demonstrated in Figure 2. In examining all variables in the two-way sensitivity analysis, the range of the test threshold, was 1.9%-15.6%, dominated by the uncertainty in the probability of death from initial missed SAH. Other significant drivers of model variation included: probability of aneurysmal versus non-aneurysmal SAH after negative head CT, probability of long-term morbidity from initial missed SAH, and probability of renal failure requiring RRT from CIN. In the probabilistic sensitivity analysis the mean testing threshold was 4.3% (95%CI, 1.4-9.3).

Discussion

The approach to diagnostic testing for specific conditions can be examined and potentially improved by consideration of a threshold for testing at which potential harms and benefits of testing are equal. Within the context of shared decision making, this information can be used to guide both physicians and patients. In the setting of potentially deadly conditions, however, the utility of this threshold is often overshadowed by barriers including defensive practice due to medico legal or professional concerns, patient expectation, poor communication, and a focus on diagnostic certainty. We are unaware of prior published literature estimating a testing threshold for lumbar puncture in the setting of potential SAH after a negative CT.

Our decision analysis calculations suggest that a reasonable test threshold for performance of lumbar puncture for the detection of subarachnoid hemorrhage in neurologically normal patients with headache and a negative non-contrast head CT is approximately 4.3%, with a range of 1.4% to 9.3% in a robust PSA analysis. These findings contrast with common practice and classical teaching both of which tend to focus on the potential benefits of diagnosis without explicit consideration of harms arising from testing.³²

The testing threshold after negative imaging in our analysis was raised by a number of factors. First, in relevant large studies subjects diagnosed with SAH by lumbar puncture appear to have mostly non-aneurysmal SAH or false positive lumbar punctures.^{9,14} Because non-aneurysmal atraumatic SAH is associated with nearly universal complete recovery without therapy, detection in such cases yields no improvement in QALYs. Second, data suggest that delays in diagnosis, while undesirable and potentially dangerous, lead to morbidity or mortality in a minority (roughly 10%) of missed aneurysmal subarachnoid hemorrhage.^{10,33-35} Third, lumbar puncture includes small but real risks of infection and injury.⁶⁻⁸ Finally, patients with positive lumbar puncture findings typically undergo angiography, incurring risks of anaphylaxis, contrast-induced nephropathy, and additional radiation exposure. Moreover, incidental aneurysms found during angiography, present in up to 7% of screening populations²¹ will commonly be interpreted as culprit lesions and undergo neurosurgical procedures²² that include considerable harm rates.³⁶

It is not surprising that the variable with the largest impact on the uncertainty of the model was the probability of death from initial missed aneurysmal SAH, as it is the primary serious outcome from not performing LP and has a wide 95%CI reported in the literature. Most other important variables were also associated with SAH outcomes. It is also of note that excluding the direct negative effects of LP (morbidity and mortality associated with infection and neurology damage) fails to lower the testing threshold below a level that would favor performing an LP after negative CT in most patients.

Using testing threshold estimates in clinical practice depends upon knowing the probability of SAH after a negative non-contrast head CT. Fortunately, an increasingly high quality database of prospective studies has begun to fill gaps that have long hampered attempts to examine this issue based on outcomes data.^{2,9,10,14} These investigations, when combined with our analysis, strongly suggest that in most patients with acute headache lumbar puncture after negative CT with newer generation scanners, especially when performed under 6 hours of symptom onset, is a more harmful than helpful strategy. This results are also supported by a recent cost-effectiveness study that demonstrated when the CT sensitivity is >99% (i.e. CT on newer generation scanner performed less 6 hours from onset of symptoms) no further testing is warranted.⁵⁸ For carefully selected patients (those with a high probability of disease (>20%) and who present late >2 days the likelihood of SAH may exceed testing thresholds in the lowest range of our intervals, suggesting that lumbar puncture may be a beneficial approach for such patients presuming the most conservative estimates for all input variables (Figure 3). Because of the declining performance of CT for SAH over time and the complicated aspect of determining a pre-CT probability of disease, we believe decision aids such as figure 3 coupled with clinical decision rules that estimate pre-CT probabilities of disease will be helpful in making shared decisions with patients under uncertainty.

Limitations

The strength of a decision analysis is dependent upon the validity of variable input and the structural assumptions of the model. In our model there are limitations based on the quality and validity of the available literature addressing each input, and the inferences that can be reasonably made from observational data. Ideally, there would be randomized trial data to inform outcome predictions based on LP and non-LP approaches following negative CT in such patients. To mitigate these uncertainties we used best available data from a rigorous literature search and review, and we employ credible intervals that offer the existing range of published data (rather than 95% confidence intervals) as a means of broadening the potential outputs from our model. In this regard we find it reassuring that varying the statistically most important inputs has a limited impact on decision-making.

As noted above, we made several assumptions regarding the structure of the model to decrease complexity. We chose not to include the short-term effects of lumbar puncture, anaphylaxis, and SAH (e.g. headache, short-term cognitive deficits) as they are extremely unlikely to contribute to any significant change in the model when compared to more serious long-term effects and death. In addition, we chose not to model cost or other diagnostic testing strategies (e.g. CT/CTA then possible LP, or MRI). Cost was not considered as there is no clear accepted standard about the cost per QALY individuals or society would be willing to pay.³⁷ The strategy of CT/CTA as an initial step is fraught with the consequences of identifying a significant portion of patients with benign headache and incidental aneurysm and previous analysis has shown this strategy to be less effective than CT/LP.³⁸ A strategy incorporating MRI was not examined because of its reduced availability in acute care settings.³⁹

Conclusion

In conclusion, our data suggest an explicit threshold approach to lumbar puncture testing for neurologically normal, CT-negative acute headache patients. Our decision analysis calculations suggest this threshold to be approximately 4.3%, with a range of 1.4% to 9.3%. In light of these data, and considering the low probability of aneurysmal SAH after a negative CT, classical teaching and current guidelines addressing testing for subarachnoid hemorrhage should be revisited.

Table 1. List of input variables for Decision Analytic Model							
Variable	Value, %	Range for Sensitivity Analysis, %	Category in Model	Source			
LP sensitivity for SAH	100	94-100	Bayesian	14,40			
LP specificity for SAH	67	63-71	Bayesian	14,40			
CTA sensitivity for Aneurysm	98	97-99	Bayesian	22,41			
CTA specificity for Aneurysm	100	97-100	Bayesian	22,41			
Probability of LP long-term morbidity	0.1	0-0.2	Probability	7			
Probability of Death from LP	0.02	0-0.1	Probability	8,24,42			
Probability of ARF requiring RRT secondary to CIN	.1	0-1	Probability	25,28			
Probability of death from ARF requiring RRT secondary to CIN	35.4	20-100	Probability	26,27			
Probability of Death from Surgery for asymptomatic aneurysm	2.5	0.8-3.2	Probability	36,43			
Probability of Long-Term Morbidity from Surgery for asymptomatic aneurysm	9.2	8.1-10.4	Probability	36,43			
Probability of death from Anaphylaxis	0.0021	0.0001-0.027	Probability	29,44			
Probability of aneurysmal vs. non-aneurysmal SAH (after negative CT) [¥]	20	10-50	Probability	20			
Probability of incidental aneurysm	2	0.4-6	Probability	21,45,46			
Probability of death SAH non-aneurysmal	2.6	0.7-9.0	Probability	47			
Probability of long-term morbidity non- aneurysmal SAH	0	0-4.8	Probability	47			
Probability of Death from SAH (treated/initial correct diagnosis)	5	2-9	Probability	10,48,49			
Probability of long-term morbidity (treated/initial correct diagnosis)	31	24-38	Probability	10,48,49			
Probability of Death from initial missed aneurysmal SAH	19	9-35	Probability	10,48,49			
Probability of long-term morbidity from initial missed aneurysmal SAH	31	17-49	Probability	10,48,49			
Annualized long-term mortality rate for SAH morbidity patients*	5	0-10	Markov	50,51			
Annual cancer rate from CT (head)	.00035	0001	Markov	52-54			
Annual mortality from cancer (head)	13%	5-50%	Markov	55,56			
Annual remission (without symptoms) from cancer	5%	1-10%	Markov	56			

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Patient Age (years)	45	18-70	Markov	9		
* Calculated from the lifetime attributable risk (LAR) of cancer incidence using the linear no-threshold model from the BEIR						

VII report¹³ and division by the LAR of cancer incidence by the number of cycles in the model to arrive at a per year risk of cancer.

¥ Conservative estimate based on available data using positive LP definition of 500 rbcs/hpf

Multiple data sources were combined where possible as a weighted averages based on study sample size.⁵⁷

Terms: LP(lumbar puncture), SAH(subarachnoid hemorrhage), ARF (acute renal failure), RRT (renal replacement therapy), CIN (contrast-induced nephropathy)

Table 2. List of othinty values for Decision Analytic Model							
Variable	Value, %	Range for Sensitivity Analysis, %	Category in Model	Source			
Utility for ARF requiring RRT	0.84	0.7-0.9	Utility	17			
Utility of LP morbidity (paraparesis)	0.7	0.6-0.9	Utility	17			
Utility of Long Term morbidity from SAH	0.76	0.6-0.9	Utility	17			
Utility of combined cancer and ARF requiring RRT*	0.67	0.49-0.81	Utility	Calculated			
Utility of combined cancer and LP morbidity*	0.56	0.49-0.81	Utility	Calculated			
Utility of combined cancer and SAH morbidity*	0.56	0.42-0.81	Utility	Calculated			
Utility of combined cancer, ARF RRT, LP morbidity*	0.47	0.29-0.73	Utility	Calculated			
Utility of combined cancer, ARF RRT, SAH Morbidity*	0.51	0.29-0.73	Utility	Calculated			
Utility of combined cancer, ARF RRT, LP and SAH Morbidity*	0.35	0.18-0.66	Utility	Calculated			
Utility of combined cancer, LP and SAH morbidity*	0.43	0.25-0.73	Utility	Calculated			
Utility of combined ARF RRT, LP and SAH*	0.45	0.25-0.73	Utility	Calculated			
Utility of combined LP and SAH morbidity*	0.53	0.36-0.81	Utility	Calculated			
Discount rate	3%		Markov	18			
* Combined utility values formed from multiplying in	dividual value	s ¹⁹		1			

Captions for Figures

Figure 1A&B

Representative components of the decision tree on whether to perform lumbar puncture after a negative non-contrast head CT. Figure 1A represents the base of tree with initial decision node (square) and subsequent downstream chance, or probability nodes (circles), and terminal nodes (triangles). Breaks in lines represent further aspects of the decision tree, part of which is demonstrated in figure 1B with Markov nodes (circles with "M").

Figure 2

Tornado diagram of two-way sensitivity analyses of variables in model and their effect on the testing threshold.

Figure 3

Conceptual model showing interaction of pre-CT probability of disease, time from onset of headache, and sensitivity of CT for subarachnoid hemorrhage. Assumed linear decrease in CT sensitivity of 5% every 12 hours, and constant specificity of 99%. Lower 95% bound (1.4) and mean value (4.3) of testing threshold are displayed. The intersection of the threshold and probability lines represent time points before which, according to our analysis, performing an LP causes more harm than good for a given pre-test probability of disease.

References :

1. Edlow JA. Diagnosis of subarachnoid hemorrhage in the emergency department. Emerg Med Clin North Am 2003;21:73-87.

2. Perry JJ, Stiell IG, Wells GA, et al. Attitudes and judgment of emergency physicians in the management of patients with acute headache. Academic emergency medicine : official journal of the Society for Academic Emergency Medicine 2005;12:33-7.

3. Connolly ES, Jr., Rabinstein AA, Carhuapoma JR, et al. Guidelines for the management of aneurysmal subarachnoid hemorrhage: a guideline for healthcare professionals from the American Heart Association/american Stroke Association. Stroke; a journal of cerebral circulation 2012;43:1711-37.

4. Edlow JA, Panagos PD, Godwin SA, Thomas TL, Decker WW. Clinical policy: Critical issues in the evaluation and management of adult patients presenting to the emergency department with acute headache. Journal of emergency nursing: JEN : official publication of the Emergency Department Nurses Association 2009;35:e43-71.

5. Sayer D, Bloom B, Fernando K, et al. An Observational Study of 2,248 Patients Presenting With Headache, Suggestive of Subarachnoid Hemorrhage, Who Received Lumbar Punctures Following Normal Computed Tomography of the Head. Academic emergency medicine : official journal of the Society for Academic Emergency Medicine 2015;22:1267-73.

Baer ET. Post-dural puncture bacterial meningitis. Anesthesiology 2006;105:381-93.

7. Dahlgren N, Tornebrandt K. Neurological complications after anaesthesia. A followup of 18,000 spinal and epidural anaesthetics performed over three years. Acta Anaesthesiol Scand 1995;39:872-80.

8. Thigpen MC, Whitney CG, Messonnier NE, et al. Bacterial meningitis in the United States, 1998-2007. The New England journal of medicine 2011;364:2016-25.

9. Perry JJ, Stiell IG, Sivilotti ML, et al. Sensitivity of computed tomography performed within six hours of onset of headache for diagnosis of subarachnoid haemorrhage: prospective cohort study. Bmj 2011;343:d4277.

6.

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10. Kowalski RG, Claassen J, Kreiter KT, et al. Initial misdiagnosis and outcome after subarachnoid hemorrhage. Jama 2004;291:866-9.

11. Kassirer JP, Moskowitz AJ, Lau J, Pauker SG. Decision analysis: a progress report. Ann Intern Med 1987;106:275-91.

12. Pauker SG, Kassirer JP. The threshold approach to clinical decision making. The New England journal of medicine 1980;302:1109-17.

13. Weinstein MC, O'Brien B, Hornberger J, et al. Principles of good practice for decision analytic modeling in health-care evaluation: report of the ISPOR Task Force on Good Research Practices--Modeling Studies. Value Health 2003;6:9-17.

14. Perry JJ, Spacek A, Forbes M, et al. Is the combination of negative computed tomography result and negative lumbar puncture result sufficient to rule out subarachnoid hemorrhage? Annals of emergency medicine 2008;51:707-13.

15. Lin N, Cahill KS, Frerichs KU, Friedlander RM, Claus EB. Treatment of ruptured and unruptured cerebral aneurysms in the USA: a paradigm shift. J Neurointerv Surg 2012;4:182-9.

16. Edlow JA, Fisher J. Diagnosis of subarachnoid hemorrhage: time to change the guidelines? Stroke; a journal of cerebral circulation 2012;43:2031-2.

17. Tengs TO, Wallace A. One thousand health-related quality-of-life estimates. Med Care 2000;38:583-637.

18. Gold MR. Cost-effectiveness in health and medicine. New York: Oxford University Press; 1996.

19. Flanagan W, McIntosh CN, Le Petit C, Berthelot JM. Deriving utility scores for comorbid conditions: a test of the multiplicative model for combining individual condition scores. Popul Health Metr 2006;4:13.

20. Perry JJ, Alyahya B, Sivilotti MLA, et al. Differentiation between traumatic tap and aneurysmal subarachnoid hemorrhage: prospective cohort study. Bmj-Brit Med J 2015;350.

21. Li MH, Chen SW, Li YD, et al. Prevalence of unruptured cerebral aneurysms in Chinese adults aged 35 to 75 years: a cross-sectional study. Ann Intern Med 2013;159:514-21.

22. Carstairs SD, Tanen DA, Duncan TD, et al. Computed tomographic angiography for the evaluation of aneurysmal subarachnoid hemorrhage. Academic emergency medicine : official journal of the Society for Academic Emergency Medicine 2006;13:486-92.

23. Perry JJ, Symington C, Mansour M, Taljaard M, Stiell IG. Is this subarachnoid hemorrhage significant? A National Survey of Neurosurgeons. The Canadian journal of neurological sciences Le journal canadien des sciences neurologiques 2012;39:638-43.

24. Evans RW. Complications of lumbar puncture. Neurol Clin 1998;16:83-105.

25. Kooiman J, Pasha SM, Zondag W, et al. Meta-analysis: serum creatinine changes following contrast enhanced CT imaging. Eur J Radiol 2012;81:2554-61.

26. McCullough PA, Wolyn R, Rocher LL, Levin RN, O'Neill WW. Acute renal failure after coronary intervention: incidence, risk factors, and relationship to mortality. Am J Med 1997;103:368-75.

27. Mitchell AM, Jones AE, Tumlin JA, Kline JA. Prospective study of the incidence of contrast-induced nephropathy among patients evaluated for pulmonary embolism by contrast-enhanced computed tomography. Academic emergency medicine : official journal of the Society for Academic Emergency Medicine 2012;19:618-25.

28. Mitchell AM, Jones AE, Tumlin JA, Kline JA. Incidence of contrast-induced nephropathy after contrast-enhanced computed tomography in the outpatient setting. Clin J Am Soc Nephrol 2010;5:4-9.

29. Cochran ST, Bomyea K, Sayre JW. Trends in adverse events after IV administration of contrast media. AJR American journal of roentgenology 2001;176:1385-8.

30. Doubilet P, Begg CB, Weinstein MC, Braun P, McNeil BJ. Probabilistic Sensitivity Analysis Using Monte Carlo Simulation. Med Decis Making 1985;5:157-77.

31. Critchfield GC, Willard KE. Probabilistic analysis of decision trees using Monte Carlo simulation. Med Decis Making 1986;6:85-92.

32. Warner JL, Najarian RM, Tierney LM, Jr. Perspective: Uses and misuses of thresholds in diagnostic decision making. Academic medicine : journal of the Association of American Medical Colleges 2010;85:556-63.

33. Kassell NF, Kongable GL, Torner JC, Adams HP, Jr., Mazuz H. Delay in referral of patients with ruptured aneurysms to neurosurgical attention. Stroke; a journal of cerebral circulation 1985;16:587-90.

34. Vermeulen MJ, Schull MJ. Missed diagnosis of subarachnoid hemorrhage in the emergency department. Stroke; a journal of cerebral circulation 2007;38:1216-21.

35. Mayer PL, Awad IA, Todor R, et al. Misdiagnosis of symptomatic cerebral aneurysm. Prevalence and correlation with outcome at four institutions. Stroke; a journal of cerebral circulation 1996;27:1558-63.

36. Wiebers D. Unruptured intracranial aneurysms: natural history, clinical outcome, and risks of surgical and endovascular treatment. The Lancet 2003;362:103-10.

37. Neumann PJ, Cohen JT, Weinstein MC. Updating cost-effectiveness--the curious resilience of the \$50,000-per-QALY threshold. The New England journal of medicine 2014;371:796-7.

38. Ward MJ, Bonomo JB, Adeoye O, Raja AS, Pines JM. Cost-effectiveness of diagnostic strategies for evaluation of suspected subarachnoid hemorrhage in the emergency department. Academic emergency medicine : official journal of the Society for Academic Emergency Medicine 2012;19:1134-44.

39. Ginde AA, Foianini A, Renner DM, Valley M, Camargo CA, Jr. Availability and quality of computed tomography and magnetic resonance imaging equipment in U.S. emergency departments. Academic emergency medicine : official journal of the Society for Academic Emergency Medicine 2008;15:780-3.

40. Claveau D, Dankoff J. Is lumbar puncture still needed in suspected subarachnoid hemorrhage after a negative head computed tomographic scan? CJEM 2013;15:1-3.

41. Westerlaan HE, van Dijk JM, Jansen-van der Weide MC, et al. Intracranial aneurysms in patients with subarachnoid hemorrhage: CT angiography as a primary examination tool for diagnosis--systematic review and meta-analysis. Radiology 2011;258:134-45.

42. Durand ML, Calderwood SB, Weber DJ, et al. Acute bacterial meningitis in adults. A review of 493 episodes. The New England journal of medicine 1993;328:21-8.

43. Raaymakers TW, Rinkel GJ, Limburg M, Algra A. Mortality and morbidity of surgery for unruptured intracranial aneurysms: a meta-analysis. Stroke; a journal of cerebral circulation 1998;29:1531-8.

44. Katayama H, Yamaguchi K, Kozuka T, Takashima T, Seez P, Matsuura K. Adverse reactions to ionic and nonionic contrast media. A report from the Japanese Committee on the Safety of Contrast Media. Radiology 1990;175:621-8.

45. Vlak MHM, Algra A, Brandenburg R, Rinkel GJE. Prevalence of unruptured intracranial aneurysms, with emphasis on sex, age, comorbidity, country, and time period: a systematic review and meta-analysis. Lancet Neurol 2011;10:626-36.

46. Rinkel GJE, Djibuti M, Algra A, van Gijn J. Prevalence and Risk of Rupture of Intracranial Aneurysms : A Systematic Review. Stroke; a journal of cerebral circulation 1998;29:251-6.

47. Rinkel GJ, Wijdicks EF, Hasan D, et al. Outcome in patients with subarachnoid haemorrhage and negative angiography according to pattern of haemorrhage on computed tomography. Lancet 1991;338:964-8.

48. Molyneux AJ, Kerr RSC, Birks J, et al. Risk of recurrent subarachnoid haemorrhage, death, or dependence and standardised mortality ratios after clipping or coiling of an intracranial aneurysm in the International Subarachnoid Aneurysm Trial (ISAT): long-term follow-up. Lancet Neurol 2009;8:427-33.

49. Hop JW, Rinkel GJ, Algra A, van Gijn J. Case-fatality rates and functional outcome after subarachnoid hemorrhage: a systematic review. Stroke; a journal of cerebral circulation 1997;28:660-4.

50. Ronkainen A, Niskanen M, Rinne J, Koivisto T, Hernesniemi J, Vapalahti M. Evidence for excess long-term mortality after treated subarachnoid hemorrhage. Stroke; a journal of cerebral circulation 2001;32:2850-3.

51. Pyysalo L, Luostarinen T, Keski-Nisula L, Ohman J. Long-term excess mortality of patients with treated and untreated unruptured intracranial aneurysms. Journal of neurology, neurosurgery, and psychiatry 2013;84:888-92.

52. Mathews JD, Forsythe AV, Brady Z, et al. Cancer risk in 680,000 people exposed to computed tomography scans in childhood or adolescence: data linkage study of 11 million Australians. Bmj 2013;346:f2360.

53. Smith-Bindman R, Lipson J, Marcus R, et al. Radiation dose associated with common computed tomography examinations and the associated lifetime attributable risk of cancer. Archives of internal medicine 2009;169:2078-86.

54. Salibi PN, Agarwal V, Panczykowski DM, Puccio AM, Sheetz MA, Okonkwo DO. Lifetime attributable risk of cancer from CT among patients surviving severe traumatic brain injury. AJR American journal of roentgenology 2014;202:397-400.

55. Siegel R, Ma J, Zou Z, Jemal A. Cancer statistics, 2014. CA: a cancer journal for clinicians 2014;64:9-29.

56. Ostrom QT, Gittleman H, Fulop J, et al. CBTRUS Statistical Report: Primary Brain and Central Nervous System Tumors Diagnosed in the United States in 2008-2012. Neuro-oncology 2015;17 Suppl 4:iv1-iv62.

57. Saramago P, Manca A, Sutton AJ. Deriving input parameters for cost-effectiveness modeling: taxonomy of data types and approaches to their statistical synthesis. Value in health : the journal of the International Society for Pharmacoeconomics and Outcomes Research 2012;15:639-49.

58. Malhotra A, Wu X, Kalra V, et al. Cost-effectiveness Analysis of Follow-up Strategies for Thunderclap Headache Patients With Negative Noncontrast CT. Academic Emergency Medicine 2016;23:243–250









Conceptual Model of Interaction of Pre-CT Probability of Disease, Time, and CT sensitivity